

AFFIRMATIVE ACTION, FACULTY PRODUCTIVITY AND CASTE INTERACTIONS: EVIDENCE FROM ENGINEERING COLLEGES IN INDIA

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Abstract

Affirmative action programs are often criticized because of concerns that they result in lower worker productivity and efficiency losses. We study the relative productivity of workers benefiting from an aggressive affirmative action policy in a setting where hiring constraints are especially likely to bind. In India, colleges are required to reserve approximately 50 percent of faculty hires for individuals from disadvantaged caste and social class groups. We collect and analyze data from a nationally representative sample of 50 engineering and technology colleges in India, some of which randomly assign students to classrooms. We find that reservation category faculty have lower levels of education, lower professorial ranks and fewer years of experience in academia than general category faculty who are not hired through reservations. Yet, even with lower qualifications, we find no evidence that reservation category faculty provide lower quality instruction across a wide range of measures that include course grades, follow-on course grades, standardized test scores, dropout, attendance, graduate school plans, and graduation. In fact, we find that, at least for immediate effects on course grades, students taught by reservation category faculty perform slightly better than students taught by general category faculty. We find no evidence of positive "teacher-like-me" effects of reservation category faculty on the relative course performance and longer-term outcomes of reservation category students. Furthermore, even in the face of potential discrimination and resentment against faculty hiring quotas, general category students perform slightly better in classrooms taught by reservation category faculty than general category faculty. The findings have implications for the heated debates over affirmative action programs found in many countries around the world and in India which is now the largest country in the world.

Keywords: Affirmative action, caste, reservation, student-faculty diversity gap, worker productivity, teacher-like-me effects, instructional quality, inequality, STEM

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1 Introduction

Organizations around the world are attempting to increase the diversity of their workforces through affirmative action programs (Fryer & Loury, 2013; Sowell, 2008). Recently, for example, large tech companies have pledged support for affirmative action programs in college admissions to help them diversify their highly educated workforce (for which they have been criticized).¹ The primary goals of affirmative action programs are to counter the effects of past discrimination and reduce economic, social and political inequality. Government departments, health care and educational institutions, and law enforcement agencies have the added goal of closer representing their workforces to the populations they serve because of the potential for positive spillovers, especially for disadvantaged groups. The potential benefits of affirmative action programs are considered so important to counteract historically ingrained discrimination that they are even included in national and state constitutions.²

Opponents of affirmative action programs often argue that workers hired through such programs, especially those that invoke quotas, have lower qualifications and are accordingly less productive (i.e. there is an equity vs efficiency tradeoff). Lower qualifications among workers targeted by affirmative action, however, do not necessarily imply lower worker productivity. For example, if workers targeted by affirmative action face discrimination in the private sector but not the public sector, then higher ability workers may sort into public sector jobs while lower ability workers may sort into private sector jobs. In this type of situation, the average productivity of targeted workers in the public sector may actually be higher than their non-targeted colleagues in the public sector. Additionally, in firms that would otherwise discriminate but instead adopt affirmative action policies, workers hired through the policies may be more qualified and productive because they no longer face discrimination (Holzer & Neumark, 1999). In fact, a sparse literature finds “clear evidence of weaker credentials but more limited evidence of weaker labor market performance among the beneficiaries of affirmative action” (pg. 474, Holzer & Neumark, 2006).

Colleges, in general, are in the unique and interesting position of increasing diversity of not only their faculty workforce, but also their student (consumer) base. In this context, an additional commonly made argument for increasing faculty diversity through affirmative action programs is to improve the performance of college students from historically disadvantaged, underrepresented, or discriminated against groups (CCCCO, 2020; CPRHE, 2018; UCOP, 2018). These faculty might serve as role models, decrease

¹In the recent Supreme Court case against Harvard University and the University of North Carolina over affirmative action in college admissions, more than 70 major corporations from a broad range of sectors signed a brief in support of continuing affirmative action programs in admissions (*Student for Fair Admissions, Inc. v. President and Fellows of Harvard College*, 2022). The Supreme Court, however, ruled on June 29, 2023 that colleges can no longer take race into consideration when granting admission offers.

²In India’s Constitution, for example, approximately half of the positions in political bodies, various forms of employment and promotion, as well as education admissions, are reserved for disadvantaged caste and social class groups (Article 15, CoI, 1948).

the likelihood of “stereotype threat” and discrimination against minority students, increase exposure to instructors with similar cultures and languages, and contribute to a sense of belonging at the university and major (Bettinger & Long, 2005; Birdsall, Gershenson, & Zuniga, 2020; Dee, 2005; Fairlie, Hoffmann, & Oreopoulos, 2014). The evidence on this important question using objective measures of productivity and productivity differences estimated without bias is especially limited.

In this paper, we examine the relative productivity of workers benefiting from an aggressive affirmative action policy in a setting where constraints on hiring a diverse qualified workforce are likely to bind. Specifically, we examine the reservation policy in colleges in India which set strict quotas on hiring half of faculty and admitting half of students from disadvantaged caste and social class groups. We first examine whether college faculty hired through quotas (“reservation category faculty”) have lower observable credentials or qualifications. We then examine heterogeneity in instructional productivity along the dimension of student’s reservation status which is related to large differences in educational attainment. We test whether reservation category faculty particularly improve the performance of reservation category students (i.e. “teacher-like-me” effects), and the related question of whether general category students perform worse (in absolute terms) in classes taught by reservation category faculty because of possible discrimination and resentment towards these faculty who are hired through quotas.

We explore these questions using a novel, large, and nationally representative dataset that we collected on faculty and undergraduate students at 50 engineering and technology colleges in India. Most of the analyses focus on a subset of these colleges that randomly assign students to classrooms. We collect and analyze a comprehensive set of measures of faculty productivity including effects on immediate course grades, follow-on course grades, test scores in basic academic skills (i.e. math and physics knowledge), electrical engineering (EE) and computer science (CS), dropouts, expected graduation with a degree and additional longer-term student outcomes, as well as faculty research productivity such as publications, grants received, and administrative activities.

Estimating the effects of being taught by reservation category faculty on student performance, however, is usually fraught with issues of potential selection bias. General category students who have more animosity or believe that they cannot learn as well from reservation category faculty might avoid classes taught by those faculty. Reservation category students may sort into classes taught by reservation category faculty, and in particular it might be the reservation category students who value those interactions the most. Both of these types of sorting by students potentially contaminate comparisons between reservation and general category faculty teaching productivity.

To address these threats to identification, we analyze data from the engineering colleges that randomly

assign students to faculty-taught sections within courses (hereafter "classrooms"). Random assignment of students to classes does not typically occur in higher education with only a few exceptions.³ Another important feature in these colleges is that student marks are given at the course level and through end-of-semester standardized exams administered and graded by a higher-level university system that includes many colleges (referred to as the "university" in the setting of these colleges) instead of assessments or evaluations by individual faculty. This grading policy rules out the possibility, for example, that reservation category faculty favorably treat reservation category students through giving higher course marks. Also, course content is standardized, and professors use a similar syllabus to that prescribed by the All India Council for Technical Education (AICTE, 2018). Random assignment in this setting also allows us to directly estimate the effects of reservation category faculty on general category students, removing the reliance on difference-in-difference estimates that use the base or majority group as a comparison group (e.g. Egalite, Kisida, & Winters, 2015; Fairlie, Hoffmann, & Oreopoulos, 2014; Gershenson, Holt, & Papageorge, 2016).⁴ We are interested in not only the relative effect of reservation category faculty on reservation category students but we are also interested in the absolute and separate effects of reservation category faculty on general category students because of potential animosity and discrimination.

We find that reservation category faculty have lower professorial ranks, fewer years of experience, and lower educational credentials than general category faculty. However, these lower observable qualifications do not translate into lower quality teaching. We find that reservation category faculty actually teach slightly better than general category faculty as measured by course grades; students taught by reservation category faculty have a higher percentile rank for a given course, with the magnitude of difference varying between 1.3 to 1.5 percentile points. The results are statistically significant, and robust to whether or not we control for various sets of faculty characteristics as controls, student fixed effects, and course fixed effects. Reservation category faculty do not put more time into teaching, measured along a range of dimensions, and thus do not provide more, but lower-quality, instruction to students. Consistent with the findings for immediate effects on course grades, we do not find evidence of negative reservation category faculty effects on longer-term outcomes such as follow-on course grades, standardized test scores after 2 and 4 years, course attendance, dropouts, expected graduation with a degree, and graduate school plans. We also do not find that research productivity and administrative service are lower among reservation category faculty than general category faculty.⁵ Taken together, the findings are consistent with discrim-

³Random assignment takes place at the U.S. Air Force Academy that provides undergraduate education for officers in the U.S. Air Force (Carrell, Page, & West, 2010). A relatively new literature uses random assignment of registration priorities and discontinuities in wait lists to provide exogenous variation in the level of course choice among college students (Kurlaender, Jackson, Howell, & Grodsky, 2014; Robles, Gross, & Fairlie, 2021).

⁴We further build on the identification provided by random assignment by estimating difference-in-difference regressions for "teacher-like-me" effects by including student, faculty, and classroom fixed effects.

⁵Engineering colleges in India have not traditionally placed an emphasis on research productivity among their faculty (similar to the typical or representative college in the U.S.). The primary basis for promotions and evaluations is experience and degree qualifications (AICTE 2010).

ination in the private labor market pushing high-ability disadvantaged caste and social class workers into academic jobs which are covered by affirmative action.

Focusing on heterogeneity in teaching productivity by student type, we do not find evidence of “teacher-like-me” effects. There is no statistically significant difference between the performance of reservation category students taught by reservation category faculty, and reservation category students taught by general category faculty. These results hold for both course grades and longer-term outcomes such as follow-on course grades, test scores, course attendance, dropout, and expected graduation with degree. We also find that even in the face of resentment and possible discrimination, general category students obtain slightly better grades (in absolute terms) in classrooms taught by reservation category faculty than general category faculty. Taken together, these findings have implications for the heated debates over affirmative action programs in many countries around the world.

Colleges in India provide an important testing ground for understanding the relative productivity of workers hired through affirmative action, in a setting where a large-scale affirmative action program for students contributes a pool of candidates who are potentially equally productive, but may not have had the requisite minimum qualifications required to apply for such jobs had it not been for affirmative action in student admissions. India is now the largest country in the world, and has the most aggressive affirmative action program in higher education in the world.⁶ Being qualified to teach at the college level is a rare skill in India, where less than 6 percent of the prime-age population has at least a Master’s degree (the minimum qualification required to teach at engineering and technology colleges) and less than 2 percent of the reservation category population has a Master’s degree (see [Table A1](#)). There are widely stated concerns about heterogeneity in faculty quality, as well as shortages of qualified faculty to teach in engineering and technology colleges ([The Hindu, 2021](#); [The Indian Express, 2017, 2018, 2021](#)).⁷ On the other hand, there is considerable discrimination in the private labor market against workers of disadvantaged caste and social class (see, for example, [Wired, 2022](#)). The Indian IT industry, in particular, has been criticized for not expanding their pool of workers to include disadvantaged caste and social class groups ([Madheswaran & Attewell, 2007](#); [Shukla, 2022](#); [Upadhy, 2007](#)).⁸ Moreover, the scale of the reservation program is immense: engineering and technology colleges employ roughly a quarter of

⁶Approximately half of faculty and student positions are reserved for the Scheduled Castes (SCs), Scheduled Tribes (STs), and Other Backward Classes (OBCs) based on their representation in the population. The Scheduled Castes (SCs) are based on the historically based caste system, the Scheduled Tribes (STs) are based on indigenous tribal membership, and the Other Backward Classes (OBCs) are based on social and educational disadvantage. In contrast, for example, in the largest higher-education system in the United States, the California Community College system, 51 percent of enrolled students are from underrepresented groups, but only 21 percent of tenured faculty are from the same groups ([Ed Source, 2020](#)).

⁷Reservation policies in India have faced substantial criticism and resistance ([MoHRD, GoI, 2020](#); [The New York Times, 2015](#); [The New York Times, 2022](#)).

⁸Lower-caste students are found to have lower returns to education ([Bertrand, Hanna, & Mullainathan, 2010](#); [Madheswaran & Attewell, 2007](#); [Mitra, 2019](#); [Shukla, 2022](#)).

a million faculty and roughly 4.5 million students are enrolled in these colleges (AICTE, 2023; Ministry of Education, GoI, 2020). Engineering colleges in India account for nearly 25 percent of all engineering degrees awarded each year globally (NSF, 2018).⁹ Finally, focusing on engineering and technology colleges is important because of the role that these colleges play in providing opportunities for upward economic and social mobility for lower-caste and lower social class groups. In this regard, our research speaks to the broader question of whether faculty that a college is mandated to hire through a caste-based quota are comparably as effective as faculty chosen by the college unrestrictedly, given a setting in which there is a large-scale educational affirmative action program producing candidates with the credentials required to accept such positions.

Our paper contributes to two major strands of the literature. First, we contribute to the literature on affirmative action policies from the vantage point of worker productivity and efficiency loss. Previous studies find that workers hired through affirmative action policies have lower qualifications but the evidence on worker productivity is extremely limited (Holzer & Neumark, 2006).¹⁰ We provide new evidence on affirmative action workers having similar productivity in a key dimension of their jobs.¹¹ Our analysis provides novel findings on affirmative action and faculty positions in general, and new evidence focusing on reservations and worker productivity in India. The literature is surprisingly thin. Research in India has primarily focused instead on reservation policies for student admissions and outcomes, and future labor market outcomes (Bagde, Epple, & Taylor, 2016; Bertrand, Hanna, & Mullainathan, 2010; Cassan, 2019; Shukla, 2022). Our paper is the first to take advantage of random assignment of students to classrooms to alleviate concerns over selection bias in estimating faculty productivity on immediate and long-term outcomes.

Second, we contribute to the growing literature on the interaction effects of disadvantaged teachers on disadvantaged students across all levels of education (i.e. “teacher-like-me” effects). Several previous studies focus on racial interactions and find evidence of strong positive student-teacher interactions by race at the primary and secondary school levels (Dee, 2004, 2005; Egalite, Kisida, & Winters, 2015; Ehrenberg, Goldhaber, & Brewer, 1995; Gershenson, Hart, Hyman, Lindsay, & Papageorge, 2022; Gershenson, Holt, & Papageorge, 2016; Lindsay & Hart, 2017; Tran & Gershenson, 2021) and college level (Birdsall, Gershenson, & Zuniga, 2020; Fairlie, Hoffmann, & Oreopoulos, 2014; Oliver, Fairlie, Millhauser, & Roland, 2021; Price, 2010).¹² With the exception of the studies using the 1985-1989 Tennessee STAR experiment,

⁹Scientists and engineers from India represent more than 20 percent of all foreign-born science and engineering degree holders working in the United States (NSF, 2018).

¹⁰Recent studies have focused on whether temporary affirmative action programs have long-term effects on employment of targeted groups. See Kurtulus (2016); A. R. Miller and Segal (2012); C. Miller (2017), for example.

¹¹or even slightly higher productivity- as measured by immediate course grades

¹²See, for example, (Carrell, Page, & West, 2010; Dee, 2005; Hoffmann & Oreopoulos, 2009) for studies of gender interactions.

however, these studies of racial interactions do not leverage random assignment of students to teachers, and thus rely on estimating relative effects instead of absolute effects. Furthermore, we address potential concerns over differential effects between immediate and longer-term educational outcomes finding similar results (Gershenson, Hart, Hyman, Lindsay, & Papageorge, 2022). Student-teacher interactions based on caste in India have been studied much less, and the evidence is limited to K-12 levels. These studies find both negative and positive interactions (Hanna & Linden, 2012; Karachiwalla, 2019; Rawal & Kingdon, 2010). Our study is the first to explore faculty-student interactions based on caste or race and affirmative action groups, in the context of post-secondary education. Random assignment to classrooms also allows us to study for the first time the broad question of how students from advantaged groups perform when taught by teachers from less-advantaged groups in the face of potential discrimination and resentment towards hiring quotas.¹³

2 Caste System and Reservation Policy Setting

The Indian caste system has existed since 1500 BC. There are four major hierarchical classes, or *varnas*, with each class consisting of potentially thousands of castes, or *jatis*, with their own hierarchies within each class. In addition, a large set of social groups, referred to as *Dalits*, were historically excluded from the four classes, and were considered “untouchable.” In addition to a signal of social hierarchy, caste has also been an indicator of occupational groups, with each caste historically mapped to an occupational guild. After independence from British colonial rule in 1947, the Indian government established an affirmative action system, called “reservations,” that sought to increase the representation of historically disadvantaged castes in public education, central and state government positions, and local and national politics through strict quotas. The groups for whom these reservations were put in place were the formerly “untouchable” castes (i.e, Scheduled Castes), marginalized indigenous groups (Scheduled Tribes), and, following the Mandal commission report in 1990, historically disadvantaged groups within the four *varnas* (Other Backward Classes).

Official, nationally-representative government reports showing caste disparities in educational and economic outcomes are limited. To fill this void, we analyzed microdata from the nationally representative Employment and Unemployment Survey conducted by India’s National Sample Survey (NSS) Organisation in 2011. The NSS microdata provide detailed information on reservation groups, educational attainment, labor market outcomes, and income. Appendix Table A1 reports differences between general category and reservation category population. Starting with educational attainment, we find large differences between the general category and reservation group population, with the general category on

¹³General category students in India express concerns about the quality of instruction and non-meritorious hiring of lower-caste faculty, and mention not putting as much effort into courses taught by lower-caste faculty (Deshpande, 2006; Jodhka & Newman, 2007).

average having spent close to 3 additional years in school, and high school and college graduation rates for the general category being 17.5 percentage points and 12.9 percentage points higher respectively. Employment in regular jobs is much lower among groups qualifying for reservation policies, with the general category population having a 14% higher regular employment rate than the reservation category population. Weekly wages, conditional on regular employment, are also much lower for the reservation category population, both for the subset of the surveyed population who are college graduates and younger college graduates between ages 25 to 45 years. Monthly per capita consumption expenditure for reservation category households is also significantly lower than general category households, both in rural and urban settings.

3 Data and Classroom Assignment

3.1 Nationally Representative Sample

To study faculty productivity and faculty-student interactions we collected student, faculty and administrative data from a nationally representative sample of 50 engineering and technology colleges in India. We drew nationally representative samples of faculty and students from broadly defined computer science (CS) and electrical engineering (EE) majors, the two largest majors in engineering and technology colleges. The sample captures the typical or representative experience of college students and faculty and does not focus on only more selective research or so-called "elite" colleges in India.

The sampling procedure consisted of three main steps.¹⁴ In the first step, we identified a broad set of CS and EE majors or departments. CS and EE related departments were selected as these departments draw the highest enrollment, accounting for approximately half of the engineering and technology college enrollment in India.¹⁵ Furthermore, these departments comprise roughly one out of every four undergraduate (bachelor's degree) majors in STEM in India. The CS departments included Computer Engineering, Computer Science Engineering, Information Science and Engineering, and Information Technology departments. The EE departments included Electrical Engineering, Electronics and Communication Engineering, Electronics and Electrical Engineering, Electronics and Instrumentation Engineering, and Electronics and Telecommunications Engineering departments. In the second step, we randomly selected colleges that had these CS and EE programs. To do so, we used administrative data on (the population frame of) all colleges with CS and EE programs in the country. We also randomly selected colleges from elite and non-elite college strata. Specifically, we used simple random sampling to select 8 elite colleges

¹⁴The first phase of data collection took place from October-December 2017. The second phase of data collection took place from January-March 2019.

¹⁵Loyalka et al. (2022) calculate these estimates using administrative data with complete national coverage in India.

and probability proportional to size sampling to select 42 non-elite colleges.¹⁶ The national sample of colleges thus represent the range of elite and non-elite institutions in India. In the third step, we sampled students within CS and EE programs in the selected universities. We first randomly sampled 1 CS department and 1 EE department from each college. In each randomly sampled department, we sampled all first-year students. For all students, we create sample weights that reflect the inverse probability of being sampled at the college, department, and student levels.

Our student survey involved collecting data on the coursework completed by students at the time of taking the survey as well as the faculty who taught these courses. We then mapped this information to the data collected from surveying faculty, where we also obtained information on a faculty’s “reservation category status,” i.e, whether they belonged to the general category or one of the three reservation category groups. In addition to the student and faculty surveys at each college, we also surveyed department heads. We collected data for 20,239 students, and data for the 2,710 faculty that taught their courses.

To collect these data, we had the full support of the government (in particular, the Ministry of Human Resource Development and the AICTE)—and hence college and department administrators—to conduct the study. We also spent considerable time training a large team of enumerators that proctored the survey and assessments in person at each college. They also remained for 2-3 days at each college to make sure that students were able to participate even if they were unavailable on a particular day. As such, response rates were extremely high. Among enrolled students at the time of the baseline, approximately 95 percent participated in the baseline survey and assessments. Similarly, among enrolled students at the time of the endline or follow-up survey, approximately 95 percent participated in the endline survey and assessments.

3.2 Faculty Characteristics and Qualifications

We report new findings on faculty characteristics and qualifications, from our nationally representative sample of 50 engineering and technology colleges. Average faculty characteristics and qualifications are reported in [Table 3.1](#). Column 1 reports means, and Column 2 reports standard deviations. Granted that engineering and technology colleges follow reservation policies, 50 percent of faculty in our nationally representative sample belong to the reservation category.¹⁷ Most engineering and technology faculty in India are at the assistant professor rank (77 percent) whereas a smaller share are associate professors (13 percent) and full professors (6 percent). On average, faculty at engineering and technology colleges have 9.49 years of work experience in higher education. In terms of educational background, master’s

¹⁶Elite institutions were defined as the India Institutes of Technology (IITs), the Indian Institutes of Information Technology (IIITs), the National Institutes of Technology (NITs), and other institutions that ranked in the top 100 of the National Institutional Ranking Framework (NIRF) rankings developed by the Ministry of Human Resource Development, Government of India.

¹⁷Engineering and technology universities in India typically advertise vacancies for permanent faculty positions separately by each reservation category, in line with hiring guidelines from AICTE ([AICTE, 2019](#), pg30).

degrees are the minimum educational requirement for faculty and are the most common education level (61 percent). We did not find any faculty with lower levels of education. Fewer faculty have a completed PhD (17 percent) or a PhD in progress (19 percent). Twenty-five percent of faculty received their degree from one of the elite engineering and technology colleges in India. Thirty-four percent of faculty are female.

Table 3.1: Faculty and Student Characteristics in Engineering and Technology Colleges in India

| Attribute | Faculty | |
|--------------------------------|----------|-------|
| | Mean | SD |
| Reservation Category | 0.50 | 0.50 |
| Assistant professor | 0.77 | 0.42 |
| Associate professor | 0.13 | 0.34 |
| Professor | 0.06 | 0.23 |
| Experience (years) | 9.49 | 6.86 |
| Highest degree Master's | 0.61 | 0.49 |
| Highest degree PhD in progress | 0.19 | 0.39 |
| Highest degree PhD | 0.17 | 0.38 |
| Degree from elite college | 0.25 | 0.43 |
| Female | 0.42 | 0.49 |
| <i>N</i> | 2710 | 2710 |
| | Students | |
| | Mean | SD |
| Reservation Category | 0.56 | 0.50 |
| Female | 0.41 | 0.49 |
| Age (years) | 18.95 | 1.49 |
| Father attended college | 0.48 | 0.50 |
| Mother attended college | 0.35 | 0.48 |
| <i>N</i> | 20239 | 20239 |
| Number of colleges | 50 | |
| Number of departments | 100 | |

Note: Estimates use department-level sampling weights defined across the full national sample of surveyed colleges (50 colleges).

3.3 Student Characteristics

Table 3.1 also reports student characteristics from our nationally representative sample of 50 colleges. Approximately 56 percent of students belong to the reservation category. The mean student age is 18.95, and 41 percent of engineering students are female. Engineering students in India come from well-educated families. Roughly one half of the students have a college-educated father, and 35 percent have a college-educated mother. These levels of educational attainment are much higher than the general population as reported in Appendix Table A1, wherein we find that less than 20 percent of even the general category population graduated from college.

3.4 Colleges and Departments with Random Assignment to Classrooms

Using surveys conducted with department heads, we found that students in a subset of departments were randomly assigned to “classrooms” or sections for all courses taken during the first two years of college. These departments indicated they used a formal, computerized procedure for the random assignment. We also obtained granular course-level grade information from these departments (in 12 colleges) for all courses taken by students during the first two years.

Students enroll in courses each term in which there are typically multiple “classrooms.” Classrooms are defined as separate course sections taught by faculty during the same term to maintain small classroom sizes. For example, Electrical Engineering 101, Spring 2019 at College A is a course that might have three separate classrooms: Section A which is taught by Faculty X, Section B is taught by Faculty Y, and Section C is taught by Faculty Z. Each classroom would have roughly one third the total course enrolment for that semester. The number of classrooms for a course ranges from 1 to 15, with a median of 3 classrooms per course. Courses are distinctly defined for each college and department.

Students within a given department generally enroll in the same set of courses prescribed during the first two years of college (AICTE, 2018). Within each of these prescribed courses the random allocation of students to course sections or classrooms within department ensures that students do not self-select into classrooms with varying compositions (in terms of proportions of reservation category faculty/peers or any correlated characteristics) of faculty and classmates. Consequently, for this sample of colleges, we can estimate the causal effects of being assigned a reservation category faculty (or other faculty characteristics) on student course grades.

3.5 Course Grades

Course grades in our sampled colleges are determined by assessing student performance on traditionally administered exams. Important to this study, course grades are assigned based on end of semester exams that are conducted and graded by a higher-level entity, which in the context of colleges in India is called the “university” and is the equivalent of a university system. Thus, faculty assigned to classrooms within the same course do not have direct control over assessing student performance. Instead, a higher-level “university” agency grades the final exams for the course for which a majority of the final grade is based.¹⁸

Grades are not standardized across the colleges. Some colleges provide letter grades whereas some col-

¹⁸Our sample includes a few departments, where a proportion of the grading structure can be under the instructors’ control. But, this proportion is small and never exceeds 30 percent in our sampled colleges.

leges provide grades on a scale of 1-100. We standardize across courses and colleges by creating a ranking of all students within a course. This creates variation in course rankings across classrooms taught by different faculty. Note that course rankings by definition have mean 50 and standard deviation 28.9, because rankings follow a uniform distribution, which has a mean of $\frac{(a+b)}{2}$ and a variance of $\frac{(b-a)^2}{12}$, with $a = 0$ and $b = 100$. Most of our analyses use college-department-course (“course”) fixed effects, alleviating concerns about comparability.

3.6 Sample with Random Assignment

For the sample of 12 colleges (20 departments) where students are randomly allocated to classrooms within courses and for which we obtained course-level grades, we have 2,268 students who are enrolled in 1,277 classrooms, within 415 distinct courses, and taught by 501 different faculty.¹⁹ Each classroom is taught by only one faculty. Students assigned the same classrooms are taught by the same faculty for the entire semester. The average classroom size is 30 students and the average course size is 92 students. Our main analysis sample follows one cohort of students over their first two years of coursework.

Appendix [Table B1](#) reports faculty qualifications and student characteristics for our sample of 12 colleges that randomly assign students to classrooms. Columns 3 and 4 reports means and standard deviations. We find that 40 percent of faculty belong to the reservation category in our sample with random assignment. Most professors are at the assistant professor rank (72%), and fewer are at the associate (18%) and full (8%) professor ranks. Faculty have 9.96 average years of experience in higher education. Most faculty have a master’s degree (51%) and fewer have a PhD in progress (15%) and completed PhD (32%). Roughly one-third of faculty earned their degree from an elite college and one-third are female. These qualifications are reasonably similar to those of faculty in the national sample. The main differences are that the sample with random assignment has a lower share of reservation category faculty and female faculty, but a higher share of faculty with a completed PhD, and faculty with degrees from elite colleges. The general patterns are similar though.

The bottom panel of Appendix [Table B1](#) reports student characteristics for our sample with random assignment. We find that 54 percent of students belong to the reservation category, 44 percent are female, and the average age is 17.72 years. We find that 50 percent have a college-educated father, and 35 percent have a college-educated mother. The student characteristics for the 12-college sample are similar to those for the national sample.

¹⁹Attrition from baseline to endline for this sample was less than 4 percent.

In addition to course grades, we collected information on several longer-term outcomes. These outcomes are measured at the end of the first two years. First, we have information on scores from standardized and proctored academic skills tests that cover basic math and physics that we administered. Test scores such as these are extremely rare in analyses of higher education. We also collected information on class attendance and dropout. To further capture effects on advanced and experiential learning in engineering, we also asked all of the students about whether they planned to eventually go to graduate school and whether students report working on research with professors.

3.7 Additional Cohort of Students

We also collected data on a few longer-term educational outcomes measured at the end of the four-year program for a second cohort of students. For this cohort of students, we collected less information on educational outcomes and do not have course grades. We combined survey information with administrative information to capture major-specific test scores (computer science and electrical engineering), graduate school plans, and expected graduation with a degree for this second cohort of students, all of which are measured at the end of the four-year programs. We also have data on faculty characteristics including reservation category status for all courses taken in the first two years for each student for our sample of colleges with random assignment. This cohort includes 2289 students taught by 650 different faculty. We use this second cohort of students to study additional long-term outcomes of students by reservation category of students and faculty.

4 Econometric Methods

4.1 Instructional Quality Regression Model

To test for differences in worker productivity as measured by instructional quality between general category and reservation category faculty, we estimate several regressions for educational outcomes. The equation in which the student course grade is the dependent variable serves as the starting point for regressions for course-level, follow-on course and longer-term educational outcomes. Course grades are a good indicator of immediate instructional quality because grading is done at the course level and not classroom level, and by an independent group and not each instructor. The base regression for student grades is the following:

$$Y_{ikcf} = \alpha + \beta_1 RT_f + \gamma_2 T_f + \lambda_k + \lambda_i + \epsilon_{ikcf} \quad (4.1)$$

where Y_{ikcf} is the outcome for student i in course k , taught in classroom c by faculty f , RT_f is a dummy variable indicating the reservation category status of faculty f (equals 1 for reservation category faculty and 0 for general category faculty), T_f is a vector of teacher characteristics for faculty f , λ_k are course fixed effects (i.e. college-department-semester offerings of courses), λ_i are student fixed effects, and ϵ_{ikcf} is the error term. Classrooms are taught by only one faculty and are within courses. Since students take multiple courses over the two-year period, we include student fixed effects that capture observable and unobserved student characteristics such as the reservation status of the student, ability, aptitudes, and socioeconomic backgrounds. Consistent with random assignment of students to classrooms estimates of β_1 are not sensitive to the exclusion of student fixed effects or controlling for or not controlling for a set of student characteristics (see Appendix [Table F1](#)).²⁰

The starting specification does not control for any faculty characteristics and qualifications to address the question of whether there are any unconditional differences in instructional quality between reservation and general category faculty. The comparison is based on the end result of the reservation or affirmative action hiring policies of the colleges. These policies might lead to hiring less qualified faculty, and the estimate of β_1 from this specification captures the unconditional difference in teaching performance on account of those policies. This specification is of most interest for evaluating the relative instructional productivity of faculty hired through reservation policies directly taking into account the effects of any differences in qualifications. Another goal is to better understand differences in quality of instruction by reservation status, conditioning on faculty qualifications. We estimate a specification that controls for a set of key qualifications which include dummy variables for highest educational degree (bachelor’s, master’s, or PhD), whether they graduated from an elite engineering college, professorial rank (assistant, associate, or full professor), and years of work experience in academia. Conditioning on these qualifications provides evidence on whether any observed productivity differential between the two groups of faculty is capturing reservation status per se, or another related characteristic. Finally, an intermediate approach is to only control for differences in rank and experience because reservation policies continue to be in adjustment with many colleges only more recently fully implementing them. With these different goals in mind, we report estimates that control for different sets of faculty qualifications.

4.2 Faculty-Student Interaction Regression Models

We next examine heterogeneity in instructional productivity by student type. To test whether reservation category students perform better when taught by reservation category faculty than with general category faculty (i.e. test for “teacher-like-me” effects), we interact the reservation category status of the student

²⁰Student characteristics include reservation category status, gender, age, mother’s education level, and father’s education level.

with that of the faculty. The same model also allows us to explore whether general category students do worse with reservation category faculty than with general category faculty. Potential reasons behind this might be resentment about being taught by reservation category faculty, resulting in lower levels of effort in those courses, or learning differences on account of a mismatch in cultural, linguistic, or other backgrounds. We start with the following model:

$$Y_{ikcf} = \alpha + \beta_1 RT_f + \beta_2 RT_f \times RS_i + \gamma_2 T_f + \lambda_k + \lambda_i + \epsilon_{ikcf} \quad (4.2)$$

where RS_i is a dummy variable for the reservation category status of student i , as defined earlier. The student fixed effects λ_i subsume the stand-alone student reservation status indicator RS_i . We estimate this base specification including various sets of faculty characteristics.

When we focus on the question of “teacher-like-me” effects instead of absolute effects we can push the model further by adding faculty fixed effects λ_f , which subsumes the reservation category status indicator for the faculty RT_f and the faculty characteristics T_f . We use variation across courses for faculty to identify these fixed effects. In another specification, we can add classroom fixed effects λ_c , which in turn subsume both the course fixed effect λ_k and the faculty fixed effect λ_f . As a result, the reservation category status indicator variables RT_f and RS_i , and fixed effects λ_k and λ_f are no longer identified. The final model is specified as:

$$Y_{ikcf} = \alpha + \beta_2 RT_f \times RS_i + \lambda_i + \lambda_c + \epsilon_{ikcf} \quad (4.3)$$

In this case β_2 is identified from comparisons between reservation category and general category students in the same classroom but with different reservation status of faculty.

5 Results

5.1 Reservation Status and Faculty Qualifications

We first examine whether faculty hired through reservation policies have lower qualifications than general category faculty. Lower qualifications may, but do not necessarily, contribute to differences in quality of instruction (Hanushek, Kain, & Rivkin, 2005) between general category and reservation category faculty. Reservation category faculty candidates are in shorter supply and thus chosen from a more restricted labor pool. We explore reservation category vs. general category differences in the population using NSS microdata, as well as among faculty using the nationally representative sample of engineering and technology colleges.

First, our analysis of NSS microdata indicates that among the broader population that belongs to groups that qualify for reservation policies, individuals are much less likely to have a master’s degree (the minimum educational credential required to teach at engineering and technology colleges in India), than individuals in the general category population. As reported in Appendix [Table A1](#), less than 2 percent of the reservation category population has a master’s degree, compared with nearly 6 percent of the general category population. The percentage of the reservation category population with a master’s degree is also lower when conditioning on younger ages, high school degrees or college degrees. These findings suggest that the general labor pool meeting the minimum educational credentials for teaching at a college is smaller for the reservation category population.

Second, using our nationally representative sample of 50 colleges, we present novel findings on the question of whether faculty hired through reservation policies have lower measurable qualifications than general category faculty. There is surprisingly little evidence on this question in the existing literature and from published government reports. [Table 5.1](#) reports average faculty qualifications (educational degrees, professorial rank, and years of experience) by reservation status and the difference between the two.²¹ Reservation category faculty are 6 percentage points more likely to be assistant professors and 5 percentage points less likely to be full professors. Consistent with lower professorial ranks, reservation category faculty have about 1 year less of work experience in academia than general category faculty (relative to a base level of 10 years of experience for general category faculty).²² We also find that reservation category faculty are 7 percentage points less likely to have completed their PhDs, and 6 percentage points more likely to have a master’s degree as their highest degree, compared to general category faculty. We also find that reservation category faculty are less likely to have degrees from elite colleges. These new findings on differences in faculty qualifications indicate that reservation category faculty have lower professorial ranks, fewer years of work experience in academia, and lower education levels.²³

²¹The patterns are similar for our subsample of 12 colleges with random assignment. We discuss this comparison below when we present results for a balance check using the sample with random assignment.

²²Within professorial ranks mean years of experience are similar except for within the full professor level where reservation category faculty have less mean experience.

²³In contrast to these differences, we find similar assignment of reservation category faculty vs general category faculty to courses by term, introductory vs advanced material, and year. See Appendix [Table C1](#).

Table 5.1: Faculty Qualifications by Reservation Status at Engineering and Technology Colleges in India

| | Reservation Cat. Faculty | General Cat. Faculty | Difference |
|--------------------------------|--------------------------|----------------------|-----------------|
| Assistant professor | 0.80 | 0.74 | 0.06** (0.03) |
| Associate professor | 0.13 | 0.14 | -0.01 (0.02) |
| Professor | 0.03 | 0.08 | -0.05*** (0.01) |
| Experience (years) | 8.91 | 10.06 | -1.15** (0.49) |
| Highest degree PhD | 0.14 | 0.21 | -0.07*** (0.02) |
| Highest degree PhD in progress | 0.18 | 0.18 | 0.00 (0.03) |
| Highest degree Master's | 0.64 | 0.58 | 0.06* (0.03) |
| Degree from elite college | 0.26 | 0.23 | 0.03 (0.03) |
| Female | 0.40 | 0.44 | -0.04 (0.03) |
| <i>N</i> | 1206 | 1485 | |

Notes: Estimates use department-level sampling weights defined across the full national sample of surveyed colleges (50 colleges). The last column reports difference in group means with standard errors in parentheses. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

We next explore whether there might be differences in unobservable ability or quality between reservation and general category faculty. In particular, discrimination in the IT labor market against highly-educated reservation category workers (Upadhyaya, 2007) could limit opportunities and “push” high-ability reservation category workers into faculty positions which are covered by affirmative action policies. In this case, the average (unobservable) quality of reservation category faculty might even be higher than that of general category faculty, conditioning on working as faculty in engineering and technology colleges. Discrimination in the private labor market might alter quality differentials in non-discriminatory or affirmative action sectors of the labor market such as government or education.

To provide some descriptive evidence on this question, we estimate the differential returns to college for general category and reservation category workers using NSS microdata.²⁴ The results, presented in Table 5.2 indicate a negative and significant wage gap for workers from the disadvantaged caste and social class groups covered by reservation policies across several specifications, and after accounting for differences in education levels, age, and occupation fixed effects. We do not find evidence of a statistically significant difference between the wages of uneducated (i.e, not college graduate) reservation and general category workers after including occupation fixed effects, which is likely due to the strong mapping between caste and occupational guilds, especially for low-skilled, informal sector jobs. However, even controlling for occupations, the wage gap for college-educated workers is large for reservation category workers. Finally, we find that the wage gap between reservation and general category college graduates is significantly larger in private sector jobs, which might push qualified reserved category workers into public sector jobs with affirmative action policies. These results are consistent with the evidence provided by Madheswaran and Attewell (2007), Bertrand et al. (2010), and Mitra (2019).

²⁴Reservation-general category population differences in educational and economic outcomes are discussed above and reported in Appendix Table A1.

Table 5.2: Returns to Education by Reservation Status

| | (I) | (II) | (III) | (IV) | (V) | (VI) |
|---------------------------------------|--|---------------------|---------------------|---------------------|---------------------|---------------------|
| | Dependent Variable: ln(Weekly Wages in Rupees) | | | | | |
| College degree | 1.273*** (0.04) | 1.152*** (0.04) | 0.497*** (0.03) | 0.493*** (0.04) | 0.208*** (0.04) | 0.596*** (0.06) |
| Res. Category | -0.261*** (0.03) | -0.137*** (0.02) | -0.031* (0.02) | -0.018 (0.02) | -0.126*** (0.03) | -0.106** (0.04) |
| College degree \times Res. Category | -0.057 (0.05) | -0.140*** (0.05) | -0.152*** (0.03) | -0.154*** (0.04) | 0.002 (0.04) | -0.161** (0.07) |
| Age | | 0.045*** (0.01) | 0.042*** (0.00) | 0.032*** (0.01) | 0.075*** (0.02) | 0.046*** (0.02) |
| Age ² | | -0.000*** (0.00) | -0.000*** (0.00) | -0.000* (0.00) | -0.001*** (0.00) | -0.000** (0.00) |
| Female | | -0.519*** (0.02) | -0.418*** (0.02) | -0.394*** (0.02) | -0.410*** (0.03) | -0.234*** (0.07) |
| Urban | | 0.441*** (0.03) | 0.197*** (0.02) | 0.195*** (0.02) | 0.286*** (0.03) | 0.179*** (0.04) |
| Occupation FE | No | No | Yes | Yes | Yes | Yes |
| Age Range | 25-64 | 25-64 | 25-64 | 25-45 | 25-64 | 25-64 |
| Job Type | All | All | All | All | Public Sector | Private Sector |
| N | 56241 | 56241 | 56241 | 40856 | 17843 | 4636 |

Notes: Estimates use microdata from the 68th Round of India's National Sample Survey, and are weighted by population using NSS multipliers. The dependent variable is the log-transformation of weekly wages reported by the respondent. The sample only includes respondents reporting non-zero wages. Standard errors are clustered at the district level. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

In the end, reservation category faculty may have lower measurable educational credentials and academic ranks, but this does not imply that they are necessarily less qualified to teach students. Discrimination in the private sector might lead high-ability (along unobservable traits) reservation category workers to faculty positions.

5.2 Quality of Instruction among Reservation Category Faculty

We next explore the question of whether there are differences between the quality of instruction provided by reservation category and general category faculty. In attempting to answer this question, there are concerns about selection bias. Reservation category faculty might be assigned to different courses, and have different students choose their classes. Sorting by students and faculty, and differential sorting into courses taught by reservation category faculty potentially contaminate comparisons between reservation category and general category professors teaching the same students. We thus focus the analysis on colleges that randomly assign students to classrooms. Students typically take a fixed set of required courses over the first two years at engineering and technology colleges in India, further limiting the potential for differential selection into courses. Course fixed effects, which are constructed uniquely for each college-department-semester-course combination, account for college and department-specific factors. Student fixed effects account for observable and unobservable baseline differences in student characteristics such as ability, aptitude, and socioeconomic status.

Before turning to the regression results, we present differences in faculty characteristics by reservation

status and conduct a balance check for the random assignment of student classrooms to faculty by reservation status for our sample of colleges with random assignment. Table 5.3 reports these results. To explore potential differences between reservation and general category faculty teaching the same courses (but different classrooms) we estimate a separate regression for each faculty characteristic (i.e. row) that includes course fixed effects and a dummy variable indicating the reservation status of the faculty. Column 3 reports the coefficient estimate on this reservation category vs. general category faculty difference, and Column 4 reports the standard error. We find that reservation category faculty have lower professorial ranks, less work experience in academia, and lower education levels in our 12-college subsample, which are similar to the patterns noted above for our national sample.

Table 5.3: Faculty Differences and Balance Checks for the Sample of Colleges with Random Assignment

| Faculty characteristics | Panel A: Faculty | | | |
|-----------------------------------|-------------------|-------|-------------------|-------|
| | Mean | SD | Res.-Gen. Faculty | SE |
| Reservation Category | 0.39 | 0.49 | 1.000 | |
| Assistant professor | 0.72 | 0.45 | 0.055 | 0.061 |
| Associate professor | 0.18 | 0.38 | 0.018 | 0.043 |
| Professor | 0.08 | 0.27 | -0.071 | 0.043 |
| Experience in years | 9.96 | 6.51 | -1.391* | 0.793 |
| Highest degree is Masters | 0.51 | 0.50 | 0.147** | 0.074 |
| Highest degree is PhD | 0.32 | 0.47 | -0.133*** | 0.041 |
| Highest degree is PhD in progress | 0.15 | 0.36 | -0.023 | 0.068 |
| Degree from elite college | 0.32 | 0.47 | -0.109* | 0.063 |
| Female | 0.33 | 0.47 | -0.008 | 0.076 |
| Student characteristics | Panel B: Students | | | |
| | Mean | SD | Res.-Gen. Faculty | SE |
| Reservation Category | 0.54 | 0.50 | -0.008 | 0.010 |
| Female | 0.44 | 0.50 | -0.002 | 0.007 |
| Age | 17.72 | 0.80 | 0.001 | 0.011 |
| Father attended college | 0.50 | 0.50 | 0.005 | 0.010 |
| Mother attended college | 0.35 | 0.48 | 0.018** | 0.008 |
| Baseline academic skills score | 0.001 | 1.00 | -0.004 | 0.020 |
| JEE Main score | 68.14 | 44.33 | 0.971 | 0.920 |
| Took JEE Main | 0.67 | 0.47 | 0.004 | 0.008 |

Notes: Means and standard deviations for general category faculty characteristics are reported in Panel A. Means and standard deviations for all sampled students are reported in Panel B. The sample of colleges with random assignment (12 colleges) is used, and the unit of analysis is a student-course. The data capture 2268 students, 501 faculty, 415 courses, and 1277 classrooms. The reservation vs general category differences control for course fixed effects, and corresponding standard errors are clustered at the faculty level. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table 5.3 also reports a balance check for student characteristics. The check suggests that faculty are essentially randomly assigned students to the classrooms that they teach within a given course. Each classroom is a course-section or "classroom" within that course (e.g. Electrical Engineering 1A or Electrical Engineering 1B) and is taught by one faculty. We find no differences in student characteristics in

between classrooms taught by reservation category faculty and classrooms taught by general category faculty with the only exception that we find a slightly higher mean value for students having a college-educated mother. The difference, however, is very small. The reservation vs general category faculty differential for student's likelihood of having a college educated mother is 0.018 relative to a mean of 0.35. We have balance on the JEE scores in our sample. We also have balance on an indicator for whether students took the JEE test. As noted below, taking the JEE exam is a positive predictor of student success. We also have balance on the baseline academic skills tests that we administered. We include student fixed effects in the regressions to control for any residual imbalance in these characteristics, as well as any (observed or unobserved) student-level factors.

Table 5.4 reports estimates of Equation 4.1. Specification I only includes the faculty reservation status indicator (Res. Cat. Faculty). We find that reservation category faculty do not teach worse, and in fact teach slightly better than general category faculty. Students in classrooms taught by reservation category faculty have slightly higher grades than students in classrooms taught by general category faculty. The difference is small at 1.44 percentile ranks (scale 1-100) but is statistically significant at the 5% level. Given that the mean percentile rank is 50, this translates into a difference of 3 percent relative to the mean (or 0.05 standard deviations using the standard deviation of 28.9 as noted above).

The additional specifications reported in Table 5.4 expand the set of controls for faculty characteristics. A pure evaluation of reservation policies might stop at Specification I and not control for any differential characteristics among reservation category faculty resulting from affirmative action policies. The unadjusted coefficient on reservation category faculty on student grades incorporates the possible lower qualifications from hiring quotas. We sequentially add faculty characteristics to move from this policy focused model to one that focuses more on estimating reservation vs general category faculty differences per se. Specification II allows for reservation category faculty to be of different ranks (i.e. assistant, associate and full professor) and years of work experience in higher education. If there was a shortage of engineering faculty in the past, it is likely that engineering and technology colleges need to hire a range of professorial ranks. Thus, some colleges might need to hire reservation category (or general category) faculty at a specific rank such as associate professors. Conditioning on hiring at this level, the reservation policy binds. In any case, we find a similar coefficient on the faculty reservation status indicator variable. The coefficient implies an effect of 1.52 course grade percentile points and is statistically significant at the 5% level.

The next column (Specification III) controls for the education level of the faculty. Interestingly, having a PhD results in lower course grades for students. As shown in Table 5.3, reservation category faculty were less likely to have a PhD. However, even though controlling for this difference works to reduce the

Table 5.4: Regressions for Student Course Grades Measuring Quality of Instruction, Reservation vs. General Category Faculty

| | I | II | III | IV |
|--------------------------------|------------------|------------------|-------------------|-------------------|
| Res. Cat. Faculty | 1.44** (0.58) | 1.52** (0.59) | 1.33** (0.57) | 1.34** (0.56) |
| Associate professor | | 0.57 (0.75) | 1.25 (0.83) | 1.27 (0.82) |
| Professor | | 1.46 (0.93) | 2.97** (1.35) | 3.18** (1.32) |
| Experience in years | | -0.02 (0.05) | -0.02 (0.05) | -0.01 (0.05) |
| Highest degree PhD | | | -2.37** (1.19) | -2.55** (1.17) |
| Highest degree PhD in progress | | | -0.75 (0.80) | -0.94 (0.82) |
| Degree from elite college | | | 0.37 (0.59) | 0.31 (0.59) |
| Female | | | | 1.09* (0.57) |
| Mean | 51.19 | 51.19 | 51.19 | 51.19 |
| N | 37767 | 37716 | 37716 | 37716 |

Notes: The dependent variable is the student course grade measured as the percentile rank in the course (1-100 scale). Grades are provided at the course level and not at the faculty-taught section level. All models are run on the sample of colleges with random assignment (12 colleges), where each observation is a student-course. All models also control for student fixed effects and course fixed effects. Standard errors are clustered at the faculty level. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

coefficient on reservation category faculty, the effect is very minor, and the coefficient remains positive (1.33) and statistically significant at the 5% level. In the final specification reported in Table 5.4 we additionally control for whether the professor is female. The coefficient estimate on reservation category faculty does not change.

All of the reported regressions include student fixed effects. We also estimate regressions that control for student characteristics instead of student fixed effects. We find very similar estimates on the reservation category faculty dummy variable for all four specifications. As an additional check, we find that the results are also very similar after removing the only elite college in the 12-college sample (which only represents 4.8 percent of the total sample).

The unit of observation in the regressions is the student course-grade which implicitly places more weight on larger classrooms. To explore whether our results are partly driven by the influence of larger classrooms, we estimate regressions in which student-course observations are weighted to equalize the influence of all classroom sizes. Specifically, each student course-grade observation is weighted by the inverse of the size of the classroom. Appendix Table F2 reports the results from estimating Equation 4.1 with (inverse) class-size weights attached to each observation. We obtain similar results to those reported

in Table 5.4. The similarity of estimates is consistent with most classroom sizes being in a narrow range around 30 students and very few with more than 100 students.

The estimation results reported in Table 5.4 test for average differences between reservation category and general category faculty. To explore heterogeneity in faculty productivity, we also construct value-added measures for each (eligible) teacher in the sample, similar in spirit to Carrell and West (2010) and Figlio, Schapiro, and Soter (2015). We construct two measures; the first uses students' prior semester GPA as a baseline measure and controls for a variety of demographic and predetermined factors. The second measure uses the score from the baseline academic skills test that students took before they started the program as the baseline measure, and controlling for other demographic and pre-determined characteristics.²⁵ Appendix Figure E1 and Figure E2 compare the distributions of faculty fixed effects estimates for both reservation category and general category faculty. The difference in the distributions of the value-added measures is fully consistent with the results reported in Table 5.4. On examining the CDFs of value-added measures, at no point in the distribution do we see reservation category faculty performing significantly worse than general category faculty.

Overall, the results show consistent and robust evidence that reservation category faculty do not provide lower quality instruction to students, and in fact provide slightly higher quality instruction. The conclusion does not depend on whether we directly compare reservation category faculty to general category faculty or control for their lower professorial ranks, less work experience in higher education, and lower levels of education.

5.3 Differences in Time Spent on Teaching Activities and Teaching Practices

Do reservation category faculty devote more time to teaching, which could explain why students in their classes do better? Reservation category faculty might be of lower quality, but put more time into teaching and helping students outside of class time, resulting in similar student performance (i.e. more effort overcomes lower per quality per unit of time). Specifically, do they devote more time and effort to teaching-related activities such as advising students or preparing lessons, which in turn compensates for lower ability? To investigate this question, we run regressions for teaching-related activities focusing on the faculty reservation status indicator coefficient (Table 5.5). We examine weekly hours on advising students, course-related work, lesson planning, teaching class, and tutoring students. We continue to use the student course-grade as the unit of analysis for consistency with the quality of instruction regressions

²⁵For the first measure, student course grades are regressed on past-semester GPA, pre-determined student characteristics, teacher fixed effects, and course fixed effects. For the second measure, student course grades are regressed on the academic skills baseline test score for the first semester, past-semester GPA's for semesters 2-4, pre-determined student characteristics, teacher fixed effects, and course fixed effects. Empirical Bayesian shrinkage estimates of teacher fixed effects are reported.

and ability to control for course fixed effects and weight by the number of students taught.²⁶ We report estimates without controlling for other faculty characteristics but report estimates with these controls in the [Online Appendix](#). The estimates are similar with or without faculty controls.

Table 5.5: Regressions for Weekly Hours Spent on Various Teaching-Related Activities, Reservation vs General Category Faculty

| | Advising Students | Course-Related Work | Lesson Planning | Teaching Classes | Tutoring |
|-------------------|-------------------|---------------------|-----------------|------------------|-----------------|
| Res. Cat. Faculty | -0.44 (0.30) | -0.27 (0.41) | 0.52 (1.01) | -0.27 (1.31) | -0.14 (0.30) |
| Mean | 3.33 | 2.98 | 7.35 | 11.02 | 2.82 |
| N | 37695 | 37797 | 37797 | 37797 | 37797 |

All models are run on the sample of colleges with random assignment (12 colleges), where each observation is a student-course. All models control for student fixed effects and course fixed effects, and standard errors are clustered at the faculty level. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

The estimates reported in [Table 5.5](#) are small in magnitude and not statistically different from 0. The estimates suggest that reservation category faculty do not spend more time on course-related work or on lesson planning, or helping students outside of the classroom through advising or tutoring.²⁷

We also ask faculty a question about about weekly hours spent teaching their classes. This variable provides a useful check that reservation category faculty are not teaching for different amounts of time than general category faculty. Classrooms within courses are scheduled for the same amount of time, and thus this question provides a quality check on both reported hours worked on activities and that reservation category and general category faculty are being compared to each other for the same courses. We find no evidence that reservation category faculty spend more time teaching their courses than general category faculty.

We also surveyed faculty on their classroom-specific pedagogical practices including a set of Teaching Practices Inventory (TPI) measures based on [Wieman and Gilbert \(2014\)](#). These TPI measures provide a test of whether there are potential differences in the types of teaching practices used in classrooms. The use of active learning techniques in the classroom, for example, is a growing teaching practice and might explain instructional quality differences between reservation and general category faculty.²⁸ Estimates reported in [Table 5.6](#) do not indicate that reservation category faculty and general category faculty are

²⁶Each observation is at the student-course level to weight the results by student contact. Thus, a faculty member teaching only a few students is weighted less than a faculty member teaching hundreds of students in the sample. Standard errors are also clustered at the faculty level to account for the variation of the dependent variable being limited to the faculty level.

²⁷We also collect information on whether students received tutoring and do not find any difference based on the percentage of courses taken with reservation category faculty by students.

²⁸Studies report that using pedagogical practices such as active and collaborative learning positively impacts student performance ([Freeman et al., 2014](#); [Hoellwarth & Moelter, 2011](#); [Porter, Bailey Lee, & Simon, 2013](#)).

implementing different teaching practices.²⁹ The findings suggest that the higher instructional quality found for reservation category faculty is not due to the use of different teaching practices instead of underlying quality differences.³⁰

Table 5.6: Regressions for Use of Teaching Practices Inventory Measures, Reservation vs. General Category Faculty

| | In-class features and activities | Assignments | Feedback and testing | Collaboration |
|-------------------|-------------------------------------|-----------------|----------------------|-----------------|
| Res. Cat. Faculty | -0.28 (0.37) | -0.25 (0.31) | -0.23 (0.46) | -0.25 (0.26) |
| Mean | 9.65 | 3.55 | 8.28 | 4.20 |
| N | 38021 | 38021 | 38021 | 38021 |

All models are run on the sample of colleges with random assignment (12 colleges), where each observation is a student-course. All models control for student fixed effects and course fixed effects, and standard errors are clustered at the faculty level. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

5.4 Additional Measures of Teaching Productivity

We explore several additional measures of teaching productivity by faculty. Productivity might differ between reservation category and general category faculty, in a way that is not captured by effects on immediate educational outcomes such as course grades. Estimates of effects on course grades, for example, might capture differences in "teaching to the test" instead of learning outcomes that extend beyond that course (Carrell & West, 2010).

We first examine faculty effects on follow-on courses. An effective instructor of a course might have positive spillovers on how students do in subsequent courses in the same subject or in general. We measure follow-on courses in two ways. First, we regress student course grades on average faculty characteristics from one prior semester. Second, we measure the reservation category variable as the proportion of reservation category faculty who taught a student over all precursor courses they took in the previous semester for that specific course. In this sense, the second definition is a subset of the first definition. In both specifications, the percentage of classes taken with reservation category faculty is included when there are multiple prior courses instead of only one. Table 5.7 reports estimates. We find no evidence of a negative reservation category faculty effect on follow-on course grades.

²⁹Estimates are similar after controlling for faculty characteristics. See [Online Appendix](#).

³⁰As a robustness check, we explore whether the main results for faculty effects on student course grades are sensitive to the inclusion of measures of teaching time and teaching practices. We estimate regressions for student course grades in which we individually add the contemporaneous teaching time and teaching practices variables to the main specifications reported in Table 5.4. We find that the coefficients for the faculty reservation status indicator variable are not sensitive to the inclusion of these variables.

Table 5.7: Regressions for Follow-on Course Grades and Test Scores, Reservation vs. General Category Faculty

| | I Follow-On Grade (Semester) | II Follow-On Grade (Course) | III Academic Skills |
|-------------------|------------------------------------|-----------------------------------|------------------------|
| Res. Cat. Faculty | 0.484 (1.660) | 0.934 (0.787) | 0.009 (0.010) |
| Student controls | FE | FE | Yes |
| Mean | 51.84 | 51.67 | -0.005 |
| N | 23218 | 11743 | 1957 |

Notes: The dependent variables are (I) grade in a follow-on course based on average faculty characteristics in one prior semester, (II) grade in a follow-on course based on average faculty characteristics for *related courses* in one prior semester, and (III) standardized scores for an academic skills test administered at the end of the first two years. For Specification III, Res. Cat. faculty is the percentage of reservation category faculty who taught all prior courses taken by the student. The Res. Cat. faculty variable is rescaled to capture the effect of changing the reservation category faculty percentage by 10 percentage points (e.g. from 0.50 to 0.60). Student controls include reservation category status, gender, age, and parents' education. All models are run for the sample with random assignment (12 colleges). Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table 5.7 also reports faculty effects for academic skills tests which further capture whether students increase general engineering-related knowledge and become more effective learners in future courses. We administered and proctored our own test for academic skills (i.e. math and physics knowledge) at the end of the first two years.³¹ Baseline test scores are included as additional controls and the reservation category faculty variable is scaled so that it can be interpreted as a change in the proportion of courses taught by reservation category faculty, by 10 percentage points. Courses taken over the first two years are used to calculate the proportion of courses taught by reservation category faculty. We find no productivity differences between reservation category faculty and general category faculty in the academic skills test scores. All of the estimates reported in Table 5.7 are similar after controlling for faculty characteristics (see Online Appendix).

We examine two additional measures of faculty productivity that capture course attendance and drop outs. Table 5.8 reports estimates. In Specification I, we measure course attendance by the average daily hours attending classes (mean=6.2). We do not find evidence of a difference between reservation category and general category faculty.³² Second, we examine administrative information on dropouts by the end of the second year. Very few students drop out of engineering colleges in the first two years (mean=0.01) or in the next two years for that matter (as we show below). We also do not find any difference between reservation category and general category faculty in affecting dropout rates among students (Table 5.8,

³¹The tests were taken by a random subset (50%) of the students in the sample.

³²We collected information on whether students received tutoring and found no difference by reservation status of faculty.

Specification II).³³

Table 5.8: Regressions for Additional Educational Outcomes, Reservation vs. General Category Faculty

| | I Hours Attended | II Dropout | III Plans for Graduate School | IV Research Assistance |
|-------------------|---------------------|------------------|-------------------------------------|------------------------------|
| Res. Cat. Faculty | 0.047 (0.050) | 0.000 (0.000) | 0.022 (0.015) | -0.003 (0.005) |
| Student controls | Yes | Yes | Yes | Yes |
| Mean | 6.18 | 0.01 | 0.61 | 0.21 |
| N | 3140 | 1965 | 2156 | 2134 |

Notes: The dependent variables are (I) hours per week spent attending classes, (II) whether a student dropped out, (III) whether the student aspired to attend graduate school after their program, and (IV) whether the student assisted a professor with their research. Res. Cat. faculty is the percentage of reservation category faculty who taught courses taken by the student, and is rescaled to capture the effect of changing the reservation category faculty percentage by 10 percentage points (e.g. from 0.50 to 0.60). The coefficients from Specification II are the marginal effects from a probit model between the dropout (0/1) outcome and the listed covariates. Student controls include reservation category status, gender, age, parents' education, and academic skills baseline z-scores. All models are run for the sample with random assignment (12 colleges), where each observation is a student-test pair. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Faculty might inspire interest in graduate school and research. We next examine whether there are productivity differences on graduate school aspirations and research work opportunities. Specifications III and IV of Table 5.8 report estimates for graduate student plans and research work with faculty, respectively. We find no evidence of differential effects by reservation status. Estimates are similar after controlling for faculty characteristics (see Online Appendix).

Focusing on the first two years of the program has the advantage of capturing immediate productivity effects, the period of random assignment of students to classrooms, and rules out the possibility of estimates being confounded by dynamic accumulation effects. As part of the project, however, we collected data on a few longer-term educational outcomes measured at the end of the four-year programs for a second cohort of students. We combined survey information with administrative information to capture major-specific test scores, graduate school plans, and expected graduation with a degree. We first examine the characteristics and test for balance for this separate cohort of students (Appendix Table G1). The average characteristics of students and faculty are similar. One difference is that this cohort of students is on average two years older, which is consistent with the baseline and follow-up surveys being conducted two years later in their studies. We also find balance on all of the student characteristics. Overall, this additional cohort of students does not appear different or face different faculty characteristics than our

³³We find that no students in our sample switch majors in the first two years and only 1 student in the sample switches in the next two years.

main cohort of students for which we have course grades.

Using this cohort of students, we examine scores on tests we administered and proctored at the end of year 4 in major-specific skills, reported in Table 5.9. The proportion of classes taught by reservation category faculty is calculated over all courses taken in the first two years for each student which is when students are randomly assigned to classrooms. We find no differential effects by the reservation category faculty percentage for either endline test score (Specifications I and II). The results for electric engineering and computer science test scores measured at the end of year 4 for this second cohort of students are consistent with what we find for academic skills test scores measured at the end of year 2 for our main cohort of students.

Table 5.9: Regressions for Additional Educational Outcomes, Reservation vs. General Category Faculty Using the Second Cohort of Students

| | I EE Test (Year 4) | II CS Test (Year 4) | III Expected Graduation (Year 4) | IV Plans for Graduate School (Year 4) |
|-------------------|--------------------------|---------------------------|--|---|
| Res. Cat. Faculty | -0.029 (0.027) | 0.020 (0.037) | -0.000 (0.000) | 0.007 (0.012) |
| Student controls | Yes | Yes | Yes | Yes |
| Mean | 0.00 | -0.03 | 0.99 | 0.51 |
| N | 1060 | 510 | 2247 | 2083 |

Notes: The dependent variables are measured at the end of year 4 and are (I) standardized test score for the electrical engineering (EE) test, (II) standardized test score for the computer science (CS) test, (III) whether the student expected to graduate, and (IV) whether the student aspired for graduate school after completing their program. Res. Cat. faculty is the percentage of reservation category faculty who taught courses taken by the student, and is rescaled to capture the effect of changing the reservation category faculty percentage by 10 percentage points (e.g. from 0.50 to 0.60). Student controls include gender, age, and parents' education. The coefficients from Specification III are the marginal effects from a probit model between the expected graduation (0/1) variable and the listed covariates. All models are run on the second cohort of students for the sample with random assignment (12 colleges), where each observation is a student. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

We also examine whether students in this cohort expect to graduate with a degree at the end of year 4. We use administrative data as well as survey data to measure expected graduation. Given the proximity to finishing their degree this measure collected at the end of year 4 is likely to be an extremely accurate predictor of actual graduation with an engineering degree. We use administrative information on their status at the end of year 4. Students who are on academic leave, detained, dropped out, expelled, left the college, medical leave, or stopped paying in the system are coded as not expected to graduate. The use of this variable then builds on what we find when we examine drop outs for the main cohort of students from the beginning of year 1 to the end of year 2. We find that a very high percentage of students expect to graduate with an engineering degree (98.9 percent). We run regressions for expected graduation using this cohort of students and find no differential effect for reservation category faculty percentage on

expected graduation. We report these new results in [Table 5.9](#), Specification III.

Finally, for this cohort of students we ask the question about whether they plan on going to graduate school. Roughly half of students at the end of year 4 report planning on going to graduate school. Regressions for graduate school plans as a dependent variable do not reveal differential effects by the reservation category of faculty. We report these new estimates in [Table 5.9](#), Specification IV.

Overall, the results for the wide range of longer-term educational outcomes and using two different cohorts of engineering students are consistent with what we find for the immediate effects on course grades. We find no evidence that reservation category faculty are less productive than general category faculty.³⁴

5.5 Research Productivity

Engineering and technology colleges in India have not traditionally placed an emphasis on research productivity among their faculty (i.e. similar to the typical or representative college in the U.S. which are not research universities). Outside of the elite institutions such as the Indian Institutes of Technology (IITs), the primary basis for promotions and evaluations is a combination of experience and degree qualifications (see [AICTE, 2010](#)).³⁵ However, some emphasis has been placed recently on research productivity. We analyze whether reservation category faculty publish less than general category faculty. We focus on two measures of research productivity in terms of publishing. We examine differences between reservation and general category faculty in: (a) number of publications per year, and (b) number of international journal publications per year. The number of publications is defined as the total number of published academic international journal articles, domestic journal articles, monographs, and edited volumes.

[Table 5.10](#) reports estimates from regressing the number of publications per year on the faculty reservation status indicator variable and additional faculty characteristics. Since we are not focusing on instructional quality (where there are concerns over student sorting) we use the full 50-college sample and faculty as the unit of analysis for these regressions. We report a set of specifications that ranges from an unconditional comparison between reservation and general category faculty to a comparison that controls for the lower professorial ranks and education levels of reservation category faculty. We find no evidence that reservation category faculty publish fewer articles than general category faculty. On average, faculty at engineering and technology colleges in India produce 2.4 publications per year.

³⁴Estimates for all outcomes are not sensitive to controlling for faculty characteristics. See [Online Appendix](#).

³⁵Seniority and qualifications factor strongly into promotions. For example, an Assistant Professor with a PhD is eligible for a higher pay grade after four years of service, and one without a PhD is eligible for a higher pay grade after six years of service (see [AICTE, 2010](#)). Conditional on a vacancy being available, a candidate with more years of experience is typically granted the promotion.

The point estimate on reservation category faculty is small and precisely estimated. Controlling for the lower likelihood of having a PhD and lower likelihood of coming from an elite college among reservation category faculty does not change the result (Specification II). Our results are robust to the inclusion of all faculty characteristics and gender of the faculty (Specifications III and IV).

Table 5.10: Regressions for Number of Publications per Year, Reservation Category vs. General Category Faculty using the National Sample

| | I | II | III | IV |
|--------------------------------|-----------------|-------------------|-------------------|-------------------|
| Res. Cat. Faculty | -0.12 (0.12) | 0.07 (0.10) | 0.09 (0.10) | 0.06 (0.11) |
| Associate professor | | 0.63*** (0.18) | 0.16 (0.20) | 0.17 (0.20) |
| Professor | | 2.49*** (0.37) | 1.54*** (0.39) | 1.51*** (0.39) |
| Experience in years | | 0.07*** (0.01) | 0.05*** (0.01) | 0.05*** (0.01) |
| Highest degree PhD | | | 1.66*** (0.21) | 1.67*** (0.21) |
| Highest degree PhD in progress | | | 0.69*** (0.14) | 0.69*** (0.14) |
| Degree from elite college | | | -0.06 (0.14) | -0.06 (0.14) |
| Female | | | | -0.20** (0.10) |
| Mean | 2.4 | 2.4 | 2.4 | 2.4 |
| N | 2691 | 2685 | 2680 | 2679 |

Notes: The regressions use department-level sampling weights, and are run at the faculty level for the national sample (50 colleges). All specifications include college and department fixed effects. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Not all publications are of the same quality and may demand a different amount of effort on the part of the faculty. We attempt to mitigate the noise from publication quality by repeating our analyses using only publications in international academic journals. Table 5.11 reports the same set of specifications as those reported in Table 5.10 but using the number of publications in international journals as the dependent variable. The mean level of publications drops from 2.4 publications per year to 0.98 publications per year when including only publications in international journals. For these more rigorous and potentially more time-consuming publications we also do not find evidence that reservation category faculty are publishing less than general category faculty. The findings are not sensitive to whether faculty qualifications are included in the regressions or not.

Table 5.11: Regressions for Number of International Publications per year, Reservation vs General Category Faculty using the National Sample

| | I | II | III | IV |
|--------------------------------|-----------------|-------------------|-------------------|-------------------|
| Res. Cat. Faculty | -0.07 (0.07) | 0.03 (0.07) | 0.04 (0.06) | 0.04 (0.07) |
| Associate professor | | 0.40*** (0.11) | 0.19** (0.10) | 0.19** (0.10) |
| Professor | | 1.60*** (0.28) | 1.18*** (0.28) | 1.17*** (0.27) |
| Experience in years | | 0.02*** (0.01) | 0.01** (0.01) | 0.01** (0.01) |
| Highest degree PhD | | | 0.73*** (0.11) | 0.74*** (0.11) |
| Highest degree PhD in progress | | | 0.28*** (0.06) | 0.28*** (0.06) |
| Degree from elite college | | | -0.08 (0.07) | -0.08 (0.07) |
| Female | | | | -0.06 (0.05) |
| Mean | 0.98 | 0.98 | 0.98 | 0.98 |
| N | 2691 | 2685 | 2680 | 2679 |

Notes: The regressions use department-level sampling weights, and are run at the faculty level for the national sample (50 colleges). All specifications include college and department fixed effects. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

We also collected information on whether these papers were published in journals covered by impact factor indices. We collected information from journals covered in the Science Citation Index (SCI), Engineering Index (EI) and Social Sciences Citation Index (SSCI). Publications covered by these indices are considered top-tier in India because they have measured impact factors. The new results are reported in Appendix [Table H1](#). As expected, faculty are publishing fewer articles on average in these journals. Among the average number of publications of roughly 1 per year in international journals, an average of 0.53 articles are published in SCI, EI or SSCI journals. We find that there is no difference between reservation category and general category faculty in the number of impact factor indexed publications.

The main course grade results are also robust to the inclusion of these two measures of publications. We estimate regressions for course grades in which we individually add the contemporaneous publications outcome variables to the main specifications reported in [Table 5.4](#). We find that the reservation category faculty coefficients are not sensitive to the inclusion of the number of publications or number of international publications.

Another measure of research productivity is whether faculty members are actively obtaining funding. We collected information on whether faculty received funding from various sources such as government

agencies, private foundations, donors, or industrial partners. We find that receiving funding is not common at engineering and technology colleges in India, with only 13 percent of faculty receiving funding over the two-year period. [Table 5.12](#) reports results from regressions for funding received by faculty. We do not find evidence that reservation category faculty are less likely to obtain funding than general category faculty. Estimates are not sensitive to controlling for professorial ranks and educational levels.

Table 5.12: Regressions for Funding Received, Reservation vs. General Category Faculty using the National Sample

| | I | II | III | IV |
|--------------------------------|-----------------|-------------------|-------------------|-------------------|
| Res. Cat. Faculty | -0.01 (0.01) | -0.00 (0.01) | -0.00 (0.01) | -0.01 (0.01) |
| Associate professor | | 0.03 (0.02) | 0.01 (0.02) | 0.01 (0.02) |
| Professor | | 0.14*** (0.04) | 0.11*** (0.04) | 0.10** (0.04) |
| Experience in years | | 0.00* (0.00) | 0.00 (0.00) | 0.00 (0.00) |
| Highest degree PhD | | | 0.05** (0.02) | 0.05** (0.02) |
| Highest degree PhD in progress | | | 0.00 (0.01) | 0.00 (0.01) |
| Degree from elite college | | | -0.03** (0.01) | -0.03** (0.01) |
| Female | | | | -0.02* (0.01) |
| Mean | 0.13 | 0.13 | 0.13 | 0.13 |
| N | 2691 | 2685 | 2680 | 2679 |

Notes: The dependent variable is any research funding received at college (0/1). The regressions use department-level sampling weights, and are run at the faculty level for the national sample (50 colleges). All specifications include college and department fixed effects. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Reservation policies might directly affect government-provided grants. We checked this by separating government funding sources from private funding sources. We find that 10 percent of faculty receive government funds and 3 percent of faculty receive private funds. We estimated two separate sets of regressions and report these in new Appendix [Table H2](#) and [Table H3](#). For both funding sources, we find no difference between reservation category and general category faculty in the receipt of grants.

As a robustness check, we also estimate publication and funding regressions using the sample of colleges with random assignment and the student-course as the unit of observation ([Appendix Table H4](#)). We find similar results for the reservation category faculty coefficient. The main exception is that we find a negative and statistically significant coefficient (-0.31) for reservation category faculty in the international publications regression. Given the lack of finding a negative effect for the broader measure of publications and the narrower measure of international publications with impact factor indices we do not put too much weight on the one negative coefficient. Our preferred results from the larger nationally

representative sample consistently do not show a negative effect and the analyses of faculty productivity for publications, grants and administrative work does not need the random assignment of students to classrooms.

5.6 Service and Administrative Work

The third main job requirement of faculty is administrative service. We collected data on whether each faculty member held an administrative position in their department or at the college. Roughly one-quarter of faculty hold an administrative position. [Table 5.13](#) reports results from regressions for whether the faculty member held an administrative position at the time of the follow-up survey. For our national sample, we do not find that reservation category faculty are less likely to hold an administrative position (although we find marginal significance without controlling for faculty qualifications). Controlling for professorial rank, experience, education, and gender we find no difference in administrative positions held.

Table 5.13: Regressions for Administrative Positions Held, Reservation vs. General Category Faculty using the National Sample

| | I | II | III | IV |
|--------------------------------|------------------|-------------------|-------------------|--------------------|
| Res. Cat. Faculty | -0.05* (0.03) | -0.02 (0.03) | -0.01 (0.03) | -0.03 (0.03) |
| Associate professor | | 0.08 (0.05) | 0.07 (0.05) | 0.08* (0.05) |
| Professor | | 0.28*** (0.06) | 0.26*** (0.06) | 0.25*** (0.06) |
| Experience in years | | 0.02*** (0.00) | 0.02*** (0.00) | 0.02*** (0.00) |
| Highest degree PhD | | | 0.05 (0.04) | 0.06 (0.04) |
| Highest degree PhD in progress | | | 0.06 (0.04) | 0.06 (0.04) |
| Degree from elite college | | | -0.02 (0.03) | -0.02 (0.03) |
| Female | | | | -0.10*** (0.02) |
| Mean | 0.28 | 0.28 | 0.28 | 0.28 |
| N | 2686 | 2685 | 2680 | 2679 |

Notes: The dependent variable is administrative position held at college (0/1). The regressions use department-level sampling weights, and are run at the faculty level for the national sample (50 colleges). All specifications include college and department fixed effects. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

6 Heterogeneity in Instructional Productivity by Student Type

We examine heterogeneity in instructional productivity along the dimension of student’s reservation status. We test whether reservation category faculty particularly improve the performance of reservation

category students (i.e. “teacher-like-me” effects), and the related question of whether reservation category faculty struggle teaching general category students. General category students might perform worse (in absolute terms) in classes taught by reservation category faculty because of possible discrimination and resentment towards hiring quotas.

We first examine student differences in the data. The qualification thresholds or cutoffs in qualifying exams for university admissions are typically lower for students belonging to reservation category groups. For instance, in 2022, the range of reservation group-based differences in cutoffs for the Joint Entrance Examination (JEE) for engineering colleges varied between 21 points and 62 points on the JEE points scale ([The Indian Express, 2022](#)). We also find major differences in family background, course performance, and baseline test scores by the reservation status of students in engineering and technology colleges in India. Appendix [Table D1](#) reports estimates of student and family background characteristics from our national sample. Reservation category students are from less-educated families on average: both fathers’ and mothers’ education levels are lower. The differences in parental education are large at nearly 20 percentage points. Reservation category students, however, are similar in terms of proportion female and by age. Consistent with different admission standards, we find that reservation category students have lower JEE Main scores (only half of all aspiring students take the JEE exam; some students take local/state-based exams instead). We also find using our own academic skills test scores that reservation students have lower scores. Finally, although not reported, reservation category students have lower course grades than general category students in our data.

For the potential to reduce inequality, an important question is whether reservation category faculty have positive relative effects on reservation category students. Reservation category faculty might serve as role models, decrease the likelihood of “stereotype threats” and discrimination against minority students, increase exposure to instructors with similar cultures and languages, and contribute to a sense of belonging at the college and major ([Bettinger & Long, 2005](#); [Dee, 2005](#); [Fairlie, Hoffmann, & Oreopoulos, 2014](#)). Students can infer caste levels from the surnames of faculty. We also explore whether general category students perform worse in classes taught by reservation category faculty, potentially due to factors such as resentment towards quotas, caste discrimination, and providing less effort in classrooms taught by those faculty.

We test these two hypotheses using [Equation 4.2](#) and report estimates in [Table 6.1](#). The main reservation category faculty coefficient captures the effect for general category students. The reservation category student variable is subsumed by the student fixed effect λ_i . Note that unlike previous studies, we can identify the absolute effect on general category students because we have random assignment to classrooms. For example, in examining racial interactions in community colleges, [Fairlie, Hoffmann, and Oreopoulos](#)

(2014) focus on relative effects instead of identifying direct effects of minority faculty on non-minority students. The focus in their study is on the minority student-minority faculty interaction. Randomization allows us to directly estimate the effect on general category students. We find that general category students do slightly better in classrooms taught by reservation category faculty than in classrooms taught by general category faculty. Having a reservation category faculty increases grades by 1.5 percentiles for general category students. The estimated effect is robust to the inclusion of various faculty characteristics.

Table 6.1: Regressions for Student Course Grades Measuring Quality of Instruction, Reservation Category Faculty Interacted with Reservation Category Students

| | I | II | III | IV |
|---------------------------------------|------------------|------------------|-------------------|-------------------|
| Res. Cat. Faculty | 1.59** (0.65) | 1.68** (0.65) | 1.48** (0.65) | 1.49** (0.64) |
| Res. Cat. Faculty x Res. Cat. Student | -0.29 (0.66) | -0.29 (0.66) | -0.27 (0.66) | -0.29 (0.66) |
| Associate professor | | 0.56 (0.75) | 1.24 (0.83) | 1.26 (0.82) |
| Professor | | 1.47 (0.93) | 2.96** (1.36) | 3.17** (1.32) |
| Experience in years | | -0.02 (0.05) | -0.02 (0.05) | -0.01 (0.05) |
| Highest degree PhD | | | -2.36** (1.19) | -2.53** (1.17) |
| Highest degree PhD in progress | | | -0.75 (0.80) | -0.95 (0.82) |
| Degree from elite college | | | 0.37 (0.59) | 0.30 (0.59) |
| Female | | | | 1.09* (0.57) |
| Mean | 51.19 | 51.19 | 51.19 | 51.19 |
| N | 37718 | 37667 | 37667 | 37667 |

Notes: The dependent variable is the student course grade measured as the percentile rank in the course (1-100 scale). Grades are provided at the course level and not at the faculty-taught section level. All models are run on the sample of colleges with random assignment (12 colleges), where each observation is a student-course. All models also control for student fixed effects and course fixed effects. Standard errors are clustered at the faculty level. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Equation 4.2 also includes an interaction between reservation category faculty and students, that indicates the relative difference or extra effect for reservation category students. We find no evidence of any positive or negative differential effect of reservation category faculty on the course grade of reservation category students, relative to general category students.³⁶

We further build on the identification provided by random assignment of students to classes in two ways. First, we estimate a set of regressions that includes student fixed effects to control for unobservable student characteristics and make the comparison between reservation and general category faculty to

³⁶The results are not sensitive to the removal of student fixed effects or controls.

teaching the same students. Second, in estimating “teacher-like-me” interactions we use regression models that include classroom (i.e. specific professor taught sections of course offerings) fixed effects which use variation between reservation and general category students when assigned to the same classroom-faculty for identification. Classroom fixed effects, which are constructed uniquely for each college-department-semester-course-classroom combination, account for classroom-specific disruptions or common shocks, differences in time of day for each class, and classroom size, among other factors. Crucially, they nest faculty fixed effects, including the reservation status of the faculty. These models combine the common difference-in-difference identification strategy used in the previous literature with our use of random assignment for identification.

Focusing on the “teacher-like-me” effects we estimate [Equation 4.3](#) and report estimates in [Table 6.2](#). Specification II repeats the main specification from [Table 6.1](#) that includes course and student fixed effects and controls for the full set of faculty characteristics. Specification II includes course, student and faculty fixed effects. The inclusion of faculty fixed effects controls for additional unobserved characteristics between reservation and general category faculty that might affect the performance of all students that they teach. The reservation category student–reservation category faculty interaction captures the relative performance of reservation category students compared with general category students with the same faculty. The reservation status interaction (i.e. “teacher-like-me”) coefficient does not change with the inclusion of these faculty fixed effects.

Specification III replaces faculty fixed effects with classroom fixed effects. Classroom fixed effects subsume faculty fixed effects because each classroom is only assigned one faculty member. The inclusion of classroom fixed effects controls for additional unobserved characteristics between classrooms taught by reservation and general category faculty, that might affect the performance of all students taught in those classrooms. The reservation category student-reservation category faculty interaction captures the relative performance of reservation category students compared with general category students *in the same classrooms*. Similar to Specification II, the reservation category student-faculty interaction does not change after including these fixed effects. Even forcing the comparison to the same faculty and the same classrooms, we do not find evidence of teacher-like-me effects. Reservation category faculty teach all students slightly better but do not teach general category students relatively worse or reservation category students relatively better.

Table 6.2: Regressions for Student Course Grades Measuring Teacher-Like-Me Interactions

| | I | II | III |
|---------------------------------------|-------------------|-----------------|-----------------|
| Res. Cat. Faculty | 1.49** (0.64) | | |
| Res. Cat. Faculty x Res. Cat. Student | -0.29 (0.66) | -0.33 (0.68) | -0.32 (0.69) |
| Associate professor | 1.26 (0.82) | | |
| Professor | 3.17** (1.32) | | |
| Experience in years | -0.01 (0.05) | | |
| Highest degree PhD | -2.53** (1.17) | | |
| Highest degree PhD in progress | -0.95 (0.82) | | |
| Degree from elite college | 0.30 (0.59) | | |
| Female | 1.09* (0.57) | | |
| Mean | 51.19 | 51.19 | 51.19 |
| N | 37667 | 37667 | 37667 |

Notes: The dependent variable is the student course grade measured as the percentile rank in the course (1-100 scale). Grades are provided at the course level and not at the faculty-taught section level. All models are run on the sample of colleges with random assignment (12 colleges), where each observation is a student-course. Specification I includes course and student fixed effects, Specification II includes Course, student, and faculty fixed effects, Specification III includes student and classroom fixed effects. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

6.1 Additional Student-Faculty Interaction Regressions

We next examine whether there is evidence of heterogeneity in within-group teacher-like-me effects across reservation groups (SC, ST, or OBC). We attempt to address this issue in two key ways. First, we replace the reservation category student indicator with dummy variables for combined classes of affirmative action (SC/ST and the relatively more advantaged OBC's). General category students continue to serve as the reference group. Table 6.3 reports estimates of Equation 4.2 expanding the set of interactions between reservation category faculty and different groups of students. We continue to find a slight positive effect of reservation category faculty on course grades for all students, but no evidence of positive interactions for either of the two subgroups of reservation category students. Splitting reservation category students into more detailed groups does not alter our initial results regarding reservation category faculty instruction quality or interactions.

Table 6.3: Regressions for Student Course Grades Measuring Quality of Instruction, Reservation Category Faculty Interacted with Detailed Reservation Category Student Groups

| | I | II | III | IV |
|-----------------------------------|------------------|------------------|-------------------|-------------------|
| Res. Cat. Faculty | 1.60** (0.65) | 1.68** (0.66) | 1.48** (0.65) | 1.50** (0.64) |
| Res. Cat. Faculty x SC/ST student | -0.49 (1.47) | -0.49 (1.47) | -0.48 (1.47) | -0.47 (1.47) |
| Res. Cat. Faculty x OBC student | -0.23 (0.59) | -0.23 (0.59) | -0.21 (0.59) | -0.23 (0.59) |
| Associate professor | | 0.56 (0.75) | 1.24 (0.83) | 1.26 (0.82) |
| Professor | | 1.47 (0.93) | 2.97** (1.36) | 3.17** (1.32) |
| Experience in years | | -0.02 (0.05) | -0.02 (0.05) | -0.01 (0.05) |
| Highest degree PhD | | | -2.36** (1.19) | -2.53** (1.17) |
| Highest degree PhD in progress | | | -0.75 (0.80) | -0.94 (0.82) |
| Degree from elite college | | | 0.37 (0.59) | 0.30 (0.59) |
| Female | | | | 1.09* (0.57) |
| Mean | 51.19 | 51.19 | 51.19 | 51.19 |
| N | 37718 | 37667 | 37667 | 37667 |

Notes: The dependent variable is the student course grade measured as the percentile rank in the course (1-100 scale). Grades are provided at the course level and not at the faculty-taught section level. All models are run on the sample of colleges with random assignment (12 colleges), where each observation is a student-course. All models also control for student fixed effects and course fixed effects. Standard errors are clustered at faculty level. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Second, we check for interaction effects between faculty and students belonging to the same reservation category group. We define a match variable which takes the value of 1 if a student-teacher pair belong to the same group among disadvantaged caste and social class groups (i.e. SC student and SC faculty, ST

student and ST faculty, OBC student and OBC faculty), and 0 otherwise. Table 6.4 reports the results of a version of Equation 4.2 with this match variable. We again find a small positive main effect of being taught by reservation category faculty, and no relative gains or losses for students resulting from being matched to a faculty of the same reservation category group.

Table 6.4: Regressions for Student Course Grades Measuring Quality of Instruction, Reservation Category Faculty Interacted with Same Reservation Category Group Student

| | I | II | III | IV |
|-----------------------------------|------------------|-------------------|-------------------|-------------------|
| Student-faculty same category | -0.08 (0.54) | -0.07 (0.54) | -0.05 (0.54) | -0.05 (0.54) |
| Res. Cat. Faculty | 1.46** (0.58) | 1.54*** (0.59) | 1.34** (0.58) | 1.35** (0.57) |
| Associate professor | | 0.56 (0.75) | 1.25 (0.83) | 1.26 (0.82) |
| Professor | | 1.47 (0.93) | 2.97** (1.36) | 3.18** (1.32) |
| Experience in years | | -0.02 (0.05) | -0.02 (0.05) | -0.01 (0.05) |
| Highest degree is PhD | | | -2.36** (1.19) | -2.54** (1.17) |
| Highest degree is PhD in progress | | | -0.76 (0.80) | -0.95 (0.82) |
| Degree from elite college | | | 0.37 (0.59) | 0.30 (0.59) |
| Female | | | | 1.09* (0.57) |
| Mean | 51.19 | 51.19 | 51.19 | 51.19 |
| N | 37718 | 37667 | 37667 | 37667 |

*** $p < 0.01$; ** $p < 0.05$; * $p < 0.1$

Student-faculty same category is defined as 1 if a student and their faculty are either both SC, both ST, both OBC, or both ‘Other’, and is defined as 0 otherwise. The dependent variable is the student course grade measured as the percentile rank in the course (1-100 scale). Grades are provided at the course level and not at the faculty-taught section level. Standard errors are clustered at the faculty level. All models are run on the sample of colleges with random assignment (12 colleges), where each observation is a student-course. All models also control for student fixed effects and course fixed effects. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

We also explore interactions between reservation category faculty and additional student characteristics. First, we examine whether reservation category faculty teach students with college-educated parents better or worse than general category faculty. We find no evidence of a differential effect for students with college-educated parents. Second, we examine whether female students perform relatively better in classrooms taught by reservation category faculty. We again find no evidence of differential reservation category faculty effects for female students.

6.2 Heterogeneous Effects on Additional Educational Outcomes

We investigate whether reservation category faculty have a positive relative effect on educational outcomes for reservation category students beyond the immediate course grade. We first examine interaction effects on follow-on course grades and test scores. Reservation category faculty might inspire more interest and motivation, and improve deeper learning in engineering among reservation category students. [Table 6.5](#) reports estimates for interaction effects for follow-on course grades, and the academic skills test that we administered at the end of the first two years. For consistency with the teacher-like-me literature we report estimates with faculty controls but we find similar results without faculty controls ([Online Appendix](#)). We do not find any interaction between reservation category faculty and reservation category students. Focusing on the main effects, we also do not find evidence of a negative effect of reservation category faculty on general category students.

Table 6.5: Regressions for Follow-on Course Grades and Test Scores, Reservation Category Faculty Interacted with Reservation Category Students

| | I Follow-On Grade (Semester) | II Follow-On Grade (Course) | III Academic Skills (z-score) |
|-----------------------------|------------------------------------|-----------------------------------|-------------------------------------|
| Res. Cat. Faculty | 0.900 (2.413) | 0.704 (1.326) | 0.008 (0.014) |
| Res. Cat. Student | | | -0.278*** (0.065) |
| R.C. Faculty × R.C. Student | 0.392 (2.640) | 0.250 (1.579) | 0.001 (0.013) |
| Student controls | FE | FE | Yes |
| Faculty controls | Yes | Yes | Yes |
| Mean | 51.84 | 51.67 | -0.005 |
| N | 23191 | 11724 | 1957 |

Notes: The dependent variables are (I) grade in a follow-on course based on average faculty characteristics in one prior semester, (II) grade in a follow-on course based on average faculty characteristics for *related courses* in one prior semester, and (III) standardized scores for an academic skills test administered at the end of the first two years. For Specifications III, Res. Cat. faculty is the percentage of reservation category faculty who taught all prior courses taken by the student. The Res. Cat. variable is rescaled to capture the effect of changing the reservation category faculty percentage by 10 percentage points (e.g. from 0.50 to 0.60). Student controls include gender, age, and parents' education. Faculty controls include professor rank, experience, highest degree, elite college, and gender. All models are run for the sample with random assignment (12 colleges). Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

We also examine interaction effects for class attendance, dropout, graduate school plans, and working for professors on research projects. For these longer-term outcomes reservation category faculty might inspire interest, provide role models, and contribute to a sense of belonging to reservation category students. Estimates are reported in [Table 6.6](#). We find no evidence of positive interaction effects for these educational outcomes.

The positive effects of reservation category faculty might show up at the end of the college experiences

Table 6.6: Regressions for Additional Educational Outcomes, Reservation Category Faculty Interacted with Reservation Category Students

| | I Hours Attended | II Dropout | III Plans for Graduate School | IV Research Assistance |
|------------------------------------|---------------------|-------------------------|-------------------------------------|------------------------------|
| Res. Cat. Faculty | 0.011 (0.062) | -0.000000 (0.000000) | 0.014 (0.017) | -0.009 (0.007) |
| Res. Cat. Student | 0.355 (0.247) | 0.000008 (0.000028) | -0.034 (0.042) | 0.016 (0.031) |
| R.C. Faculty \times R.C. Student | 0.066 (0.055) | -0.000000 (0.000001) | 0.002 (0.009) | 0.006 (0.006) |
| Student controls | Yes | Yes | Yes | Yes |
| Faculty controls | Yes | Yes | Yes | Yes |
| Mean | 6.18 | 0.01 | 0.61 | 0.21 |
| N | 3140 | 1965 | 2156 | 2134 |

Notes: The dependent variables are (I) hours per week spent attending classes, (II) whether a student dropped out, (III) whether the student aspired to attend graduate school after their program, and (IV) whether the student assisted a professor with their research. Res. Cat. faculty is the percentage of reservation category faculty who taught courses taken by the student, and is rescaled to capture the effect of changing the reservation category faculty percentage by 10 percentage points (e.g. from 0.50 to 0.60). The coefficients from Specification II are the marginal effects from a probit model between the dropout (0/1) variable and the listed independent variables. Student controls include gender, age, parents' education, and academic skills baseline z-scores. Faculty controls include professor rank, experience, highest degree, elite college, and gender. All models are run for the sample with random assignment (12 colleges), where each observation is a student-test pair. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

of students. In Appendix [Table G2](#) we report estimates of reservation category faculty and student interactions for expected graduation with degree, and graduate school plans measured at the end of year 4 for our second cohort of students. We do not find teacher-like-me effects for these longer-term outcomes. We also examine interaction effects on CS and EE test scores for the second cohort of students. We find no evidence of an interaction effect for either test score.

For all of these longer-term outcomes which are measured at the end of year 2 for the main analysis sample or the end of year 4 for the second cohort of students, we find consistent results.³⁷ There is no evidence of positive or negative teacher-like-me effects on longer-term outcomes. We also do not find evidence of a negative effect of reservation category faculty on general category students.

7 Conclusion

Although the evidence is limited, affirmative action programs, especially those involving strict quotas, are often criticized because of fears that they result in lower worker productivity ([Holzer & Neumark, 2000, 2006](#)). We explore this criticism by examining the relative productivity of workers benefiting from an aggressive affirmative action policy in a setting where constraints on hiring a diverse qualified workforce are likely to bind. In India, colleges are required to reserve approximately 50 percent of faculty hires

³⁷Results without faculty controls are similar ([Online Appendix](#)).

for individuals from disadvantaged caste and social class groups to match the population. We use our nationally representative sample of 50 engineering and technology colleges in India and subset of colleges that randomly assign students to classrooms to provide novel evidence on this fundamental and understudied question about affirmative action and worker productivity. In terms of qualifications, we find that reservation category faculty have lower levels of education, lower professorial ranks and less years of work experience in academia than general category faculty. Reservation category faculty, for example, are more likely to have master's degrees and less likely to have PhDs. Yet, even with lower qualifications, we find no evidence that reservation category faculty provide lower quality instruction than general category faculty. In fact, we find that students taught by reservation category obtain slightly higher grades than students taught by general category faculty. Furthermore, even in light of potential resentment and animosity towards professors hired through reservation quotas, we find that general category students actually do slightly better (in grades) when taught by reservation category faculty. We do not find that reservation category faculty spend more time on teaching activities, and thus compensate for having lower qualifications by devoting more time to preparing and teaching classes, or advising and tutoring students.

Our estimates of differential faculty effects on longer-term educational outcomes are quite consistent across several measures. For example, we find no differential effects on follow-on course grades, and academic skills, computer science, and electrical engineering tests. The findings rule out that possibility of "teaching to the test" and suggest that reservation category faculty are not inferior at teaching higher-order engineering skills. Furthermore, we do not find lower instructional productivity as measured by longer-term outcomes such as course attendance, dropouts, expected graduation with a degree, graduate school plans, and research work with faculty. These findings are consistent across the two cohorts of students that we follow and their different stages in their studies captured.

Although teaching is the primary focus of the typical or representative college and instructional productivity has the added importance of affecting the future labor market outcomes of students, we also examine faculty's research productivity and administrative service. We do not find that reservation category faculty have different levels of research and service productivity than general category faculty.

Our results are especially compelling as we overcome traditional obstacles in establishing causality by leveraging the random assignment of students to classrooms as well as objective and accurate measures of teaching productivity (such as administrative grades, or standardized, third-party proctored test outcomes). We also focus on a large and important workforce which affects not only their own earnings but also the future earnings of students they teach. There are nearly a quarter of a million faculty, training close to 4.5 million students in engineering and technology colleges in India, with a growing number of

graduates being hired in the United States and other countries.³⁸ In this context, we find that even with an affirmative action program that has large quotas and affects a highly-educated population, the popular view should not assume that these programs result in lower worker productivity.³⁹ Potential discrimination in the broader uncovered labor market might “push” higher-ability (e.g. more enthusiastic, articulate, motivated, etc.) reservation category workers into academia. More research using careful empirical designs such as the one used here are needed to test whether affirmative action programs lead to lower worker productivity for the targeted group in other settings. Future evidence along these lines is crucial to better inform the heated debate over affirmative action programs around the world. In India, for example, reservation policies have been protested widely even invoking riots.⁴⁰ Our findings speak not only to the debate surrounding affirmative action in faculty hiring, but also to the broader system of affirmative action system in Indian education, since we are examining the relative productivity of two groups of faculty- one hired through a quota system and one hired without quotas, in a setting where an equally aggressive quota system at the student level produces candidates who otherwise may not have been able to furnish the minimum qualifications required to apply for these jobs.

Affirmative action policies often promote hiring disadvantaged and underrepresented groups with the goal of reducing inequality among the population served.⁴¹ In education, several previous studies find large, positive “teacher-like-me” effects by which teachers from underrepresented racial groups improve the academic outcomes of similar students that they teach.⁴² In our analysis of heterogeneity of instructional productivity, we do not find evidence of positive “teacher-like-me” effects of being taught by reservation category faculty on the performance of reservation category students relative to general category students. The finding is consistent across an extensive set of immediate and longer-term educational outcomes. One reason for the lack of effects is that caste discrimination might be more ingrained among students and even reservation students might associate reservation faculty as being less qualified to teach (instead of serving as a positive role model [Karachiwalla \(2019\)](#)). Another reason might be the considerable within-group heterogeneity of reservation groups. Finally, disadvantaged caste and social

³⁸One in five foreign-born science and engineering degree holders working in the United States are from India ([NSF, 2018](#)).

³⁹This is the view at the highest levels of higher education in India. Due to concerns over limiting the quality of instruction and research, The Central Educational Institutions (Reservation in Teachers’ Cadre) Act allows for “institutions of excellence, research institutions, institutions of national and strategic importance” to be exempted from reservation requirements in the Constitution ([MoE, GoI, 2019](#)).

⁴⁰See ([BBC News, 2015](#); [The New York Times, 2015](#); [The New York Times, 2022](#)). Even the specific use of disadvantaged caste and social class quotas for faculty in the elite IIT’s has been debated extensively with the 2019 Ramagopal Rao Committee Report ([MoHRD, GoI, 2020](#)) arguing to abolish reservations and the Supreme Court of India in a recent case in 2022 directing IITs to follow reservation policies. Recent evidence indicates that they are generally not following quotas in hiring faculty ([Paliwal, 2023](#)).

⁴¹Lower-caste students are underrepresented in competitive, well- paying private jobs in STEM contributing to broader caste inequality ([Deshpande, 2006](#); [Upadhya, 2007](#)).

⁴²See [Dee \(2004, 2005\)](#); [Egalite, Kisida, and Winters \(2015\)](#); [Ehrenberg, Goldhaber, and Brewer \(1995\)](#); [Gershenson, Hart, Hyman, Lindsay, and Papageorge \(2022\)](#); [Gershenson, Holt, and Papageorge \(2016\)](#) for evidence at primary and secondary school levels, and [Birdsall, Gershenson, and Zuniga \(2020\)](#); [Fairlie, Hoffmann, and Oreopoulos \(2014\)](#); [Oliver, Fairlie, Millhauser, and Roland \(2021\)](#); [Price \(2010\)](#) for evidence at the college level.

class faculty are also more prevalent at colleges because of 50 percent quotas potentially resulting in less of a role model effect. Role models might be strongest for the least represented groups among faculty. These new findings on caste interactions contribute to the scant literature which finds mixed results and focuses on K-12 education ([Karachiwalla, 2019](#); [Rawal & Kingdon, 2010](#)).

Affirmative action programs are hotly debated and facing legal challenges around the world. These programs, especially ones with quotas, are criticized because of fears that they lead to lower qualifications and preparation, lower productivity and reverse discrimination. On the other hand, proponents argue that affirmative action programs address equity concerns in employment, fight historical discrimination, and provide role models and networks for future hires.⁴³ In education there is the additional argument that hiring faculty from underrepresented groups could not only provide jobs to those groups but also could help disadvantaged and underrepresented students, both reducing inequality. The empirical evidence on both sides of this important debate, however, is limited. We provide one of the first studies of worker productivity and college student performance in the context of a strict affirmative action program in hiring and admissions. More research using careful empirical designs and the comprehensive approach taken here are needed to shed light on this multi-faceted and heated debate.

⁴³There is concern that caste discrimination has followed immigrants in host countries such as the United States leading to arguments for caste being added to protected group lists ([NBC News, 2022](#); [Equality Labs, 2018](#)). The California State University (CSU) system recently added caste to its list of protected statuses (see [CSU, 2023](#))

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Appendices

A Descriptive Statistics from NSS Micro Data

Table A1: Descriptive Statistics

| | | I General | II Reservation | III General vs Reservation |
|--|------|--------------|-------------------|----------------------------------|
| Mean Years of Schooling | Mean | 8.0 | 5.2 | 2.99*** |
| | SD | 5.2 | 4.8 | (0.03) |
| | n | 39707 | 89223 | |
| Proportion Graduating High School (%) | Mean | 29.2 | 11.7 | 17.48*** |
| | SD | 45.5 | 32.2 | (0.22) |
| | n | 39709 | 89237 | |
| Proportion Graduating College (%) | Mean | 19.4 | 6.5 | 12.88*** |
| | SD | 39.5 | 24.6 | (0.18) |
| | n | 39709 | 89237 | |
| Proportion with Master's or Higher (%) | Mean | 5.8 | 1.8 | 4.0*** |
| | SD | 23.3 | 13.2 | (0.11) |
| | n | 39709 | 89237 | |
| Proportion with Master's or Higher (%) (Age 25-50) | Mean | 6.04 | 1.98 | 4.06*** |
| | SD | 23.8 | 13.9 | (0.12) |
| | n | 31706 | 72612 | |
| Proportion with Regular Employment (%) | Mean | 30.8 | 16.9 | 13.97*** |
| | SD | 46.2 | 37.4 | (0.25) |
| | n | 39709 | 89237 | |
| Monthly Per Capita Consumption Expenditure (Rs) | Mean | 7192.8 | 5554.8 | 1638.01*** |
| | SD | 5294.5 | 4040.2 | (36.54) |
| | n | 21227 | 60066 | |
| Weekly Wages (Rupees) | Mean | 2752.2 | 1399.9 | 1363.5*** |
| | SD | 3820.9 | 1731.7 | (23.55) |
| | n | 16568 | 40135 | |
| Weekly Wages of College Graduates (Rupees) | Mean | 5747.5 | 3967.7 | 1779.7*** |
| | SD | 5733.0 | 2904.9 | (84.0) |
| | n | 5445 | 6424 | |
| Weekly Wages of College Graduates (Age 25-35) (Rupees) | Mean | 4536.3 | 3159.9 | 1376.5*** |
| | SD | 4112.6 | 2382.2 | (95.2) |
| | n | 2111 | 2764 | |

Note: Estimates are calculated using microdata from the National Sample Survey Organization's 68th Round: Employment and Unemployment Survey of 2011-12, and weighted by population using NSS multipliers. Column III reports the difference between the means in Column I (general) and column II (reservation), with the standard errors reported in parentheses. Column III reports the general category-reservation category difference in means. Monthly per capita consumption expenditure is computed at the household level. Significance levels: *** $p < 0.01$; ** $p < 0.05$; * $p < 0.1$.

B Summary Statistics: Sample of Colleges with Random Assignment

Table B1: Faculty and Student Characteristics: Sample of Colleges with Random Assignment

| Attribute | Faculty | |
|--------------------------------|----------|------|
| | Mean | SD |
| Reservation Category | 0.40 | 0.49 |
| Assistant professor | 0.72 | 0.45 |
| Associate professor | 0.18 | 0.38 |
| Professor | 0.08 | 0.27 |
| Experience (Years) | 9.96 | 6.51 |
| Highest Degree Master's | 0.51 | 0.50 |
| Highest Degree PhD in progress | 0.15 | 0.36 |
| Highest Degree PhD | 0.32 | 0.47 |
| Degree from Elite College | 0.32 | 0.47 |
| Female | 0.33 | 0.47 |
| <i>N</i> | 501 | 501 |
| | Students | |
| | Mean | SD |
| Reservation Category | 0.54 | 0.50 |
| Female | 0.44 | 0.50 |
| Age (years) | 17.72 | 0.80 |
| Father attended college | 0.50 | 0.50 |
| Mother attended college | 0.35 | 0.48 |
| <i>N</i> | 2268 | 2268 |
| Number of colleges | 12 | 12 |
| Number of departments | 20 | 20 |

C Course Assignment by Faculty Group

Table C1: Course Assignments by Faculty Reservation Category Status

| Panel A: Nationally Representative Sample | | |
|--|---|-------------------------------------|
| | Reservation Category Faculty | General Category Faculty |
| Semester 1 | 48.4% | 51.6% |
| Semester 2 | 45.0% | 55.0% |
| Semester 3 | 50.4% | 49.6% |
| Semester 4 | 53.3% | 46.7% |
| N | 95400 | 114993 |
| Panel B: Sample with Random Assignment | | |
| | Reservation Category Faculty | General Category Faculty |
| Semester 1 | 36.4% | 63.6% |
| Semester 2 | 35.4% | 64.6% |
| Semester 3 | 45.6% | 54.4% |
| Semester 4 | 42.2% | 57.8% |
| Introductory Courses | 34.4% | 65.6% |
| Advanced Courses | 46.0% | 54.0% |
| N | 14938 | 23083 |
| Panel C: Nationally Representative Sample (Second Cohort) | | |
| | Reservation Category Faculty | General Category Faculty |
| Year 1 | 43.83% | 56.17% |
| Year 2 | 46.0% | 54.0% |
| Year 3 | 46.6% | 53.4% |
| Year 4 | 44.6% | 55.4% |
| N | 172686 | 231589 |

Notes: Panel A reports the percentage of all courses (classrooms) in each semester of the first two years of the program, assigned to reservation category and general category faculty for the full sample of 50 colleges. Panel B reports the percentage of all courses (classrooms) in each semester of the first two years of the program, assigned to reservation category and general category faculty for the sample of colleges with random assignment (12 colleges), for the first cohort of students. Panel C reports the percentage of all courses (classrooms) in each semester of all four years of the program, assigned to reservation category and general category faculty for the full sample of 50 colleges, for the second cohort of students. The unit of analysis is a student-course.

D Student Differences by Reservation Category

Table D1: Reservation and General Category Student Differences in Engineering and Technology Colleges in India

| | Reservation Category Students | General Category Students | Difference | Sample size |
|--------------------------------|----------------------------------|------------------------------|-----------------|-------------|
| Female | 0.41 | 0.40 | 0.01 (0.01) | 20117 |
| Age (years) | 18.92 | 18.99 | -0.07* (0.04) | 17492 |
| Father attended college | 0.40 | 0.58 | -0.18*** (0.01) | 20062 |
| Mother attended college | 0.27 | 0.46 | -0.19*** (0.01) | 20059 |
| JEE Main score | 69.33 | 79.06 | -9.73*** (1.21) | 10259 |
| Baseline academic skills score | -0.10 | 0.12 | -0.22*** (0.02) | 8748 |
| <i>N</i> | 9619 | 10501 | | 20120 |

Notes: Estimates use department-level sampling weights defined across the full national sample of surveyed colleges (50 colleges). The last column reports difference in group means with standard errors in parentheses. JEE Main score can range between -120 (as students received a penalty for incorrect answers) and 360 . Baseline academic skills test scores are z-scores standardized across all respective test takers. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

E Value-Added Measures of Faculty Productivity

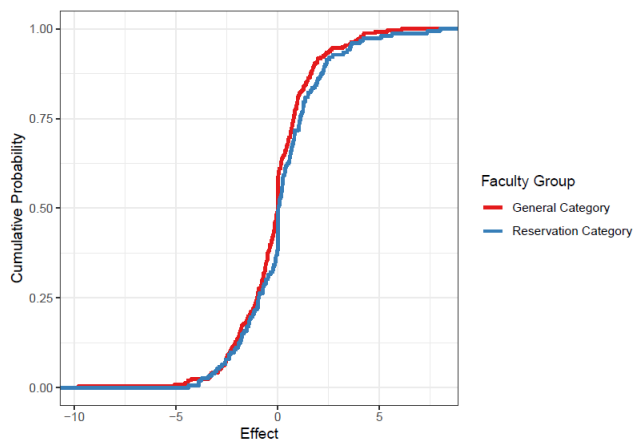


Figure E1: Distributions of Value-Added Measures of Productivity by Reservation Category: First Semester Not Included

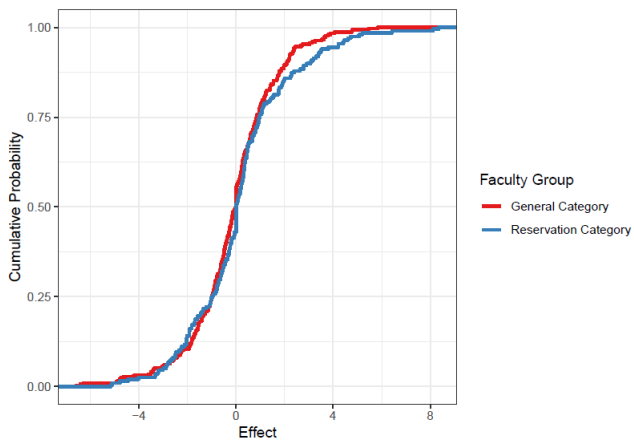


Figure E2: Distributions of Value-Added Measures of Productivity by Reservation Category: First Semester Included

F Robustness Checks

F.1 Main Results without Student Fixed Effects

Table F1: Regressions for Student Course Grades Measuring Quality of Instruction, Reservation vs. General Category Faculty: Without Student Fixed Effects

| | I | II | III |
|--------------------------------|-------------------|--------------------|-------------------|
| Res. Cat. Faculty | 1.22* (0.64) | 1.20* (0.64) | 1.34** (0.56) |
| Associate professor | 1.02 (0.88) | 1.07 (0.89) | 1.27 (0.82) |
| Professor | 4.03*** (1.38) | 4.04*** (1.46) | 3.18** (1.32) |
| Experience in years | 0.00 (0.06) | 0.02 (0.06) | -0.01 (0.05) |
| Highest degree PhD | -2.94** (1.18) | -3.29*** (1.21) | -2.55** (1.17) |
| Highest degree PhD in progress | -0.31 (1.01) | -0.52 (1.03) | -0.94 (0.82) |
| Degree from elite college | 0.06 (0.72) | 0.10 (0.71) | 0.31 (0.59) |
| Female | 1.40** (0.63) | 1.47** (0.63) | 1.09* (0.57) |
| Student characteristics | None | Main Controls | Fixed Effects |
| N | 37716 | 37716 | 37716 |

Notes: The dependent variable is the student course grade measured as the percentile rank in the course (1-100 scale). Grades are provided at the course level and not at the faculty-taught section level. Column (I) reports the results without using any student-level controls or fixed effects, column (II) uses student controls (their reservation status, age, gender, and parents' education), and column (III) includes student fixed effects, replicating the specification of [Table 5.4](#). Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

F.2 Inverse-Weighted Observations by Classroom Size

Table F2: Regressions for Student Course Grades Measuring Quality of Instruction Weighting Each Classroom the Same, Reservation vs. General Category Faculty

| | I | II | III | IV |
|-----------------------------------|----------------|-----------------|-----------------|-----------------|
| Res. Cat. Faculty | 1.52 (0.93) | 1.69* (0.93) | 1.65* (0.92) | 1.73* (0.91) |
| Associate professor | | -0.87 (1.03) | -0.67 (1.15) | -0.72 (1.14) |
| Professor | | 2.03 (1.73) | 2.49 (1.95) | 2.66 (1.93) |
| Experience in years | | 0.01 (0.07) | 0.01 (0.07) | 0.01 (0.07) |
| Highest degree is PhD | | | -0.80 (1.52) | -1.05 (1.52) |
| Highest degree is PhD in progress | | | -0.16 (1.20) | -0.31 (1.20) |
| Degree college elite | | | 0.21 (1.10) | 0.11 (1.10) |
| Female | | | | 1.41 (0.88) |
| N | 37767 | 37716 | 37716 | 37716 |

Notes: The dependent variable is the student course grade measured as the percentile rank in the course (1-100 scale). Grades are provided at the course level and not at the faculty-taught section level. Reciprocal of classroom sizes are used as regression weights. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

G Characteristics and Outcome Regressions for Second Cohort of Students

Table G1: Faculty Differences and Balance Checks for the Sample of Colleges with Random Assignment (Second Cohort of Students)

| Panel A: Faculty | | | | |
|-----------------------------------|--------|-------|-------------------|-------|
| Faculty characteristics | Mean | SD | Res.-Gen. Faculty | SE |
| Reservation Category | 0.38 | 0.48 | 1.000 | |
| Assistant professor | 0.72 | 0.45 | 0.095** | 0.048 |
| Associate professor | 0.16 | 0.37 | -0.044 | 0.040 |
| Professor | 0.08 | 0.27 | -0.034 | 0.034 |
| Experience in years | 10.05 | 6.84 | -1.035 | 0.684 |
| Highest degree is Masters | 0.54 | 0.50 | 0.005 | 0.053 |
| Highest degree is PhD | 0.25 | 0.43 | -0.054 | 0.042 |
| Highest degree is PhD in progress | 0.19 | 0.40 | 0.032 | 0.048 |
| Degree from elite college | 0.30 | 0.46 | 0.083* | 0.045 |
| Female | 0.33 | 0.47 | -0.036 | 0.053 |
| Panel B: Students | | | | |
| Student characteristics | Mean | SD | Res.-Gen. Faculty | SE |
| Reservation Category | 0.49 | 0.50 | -0.003 | 0.009 |
| Female | 0.45 | 0.50 | -0.004 | 0.008 |
| Age | 19.76 | 0.99 | 0.011 | 0.013 |
| Father attended college | 0.56 | 0.50 | 0.003 | 0.007 |
| Mother attended college | 0.39 | 0.49 | -0.002 | 0.006 |
| Baseline academic skills score | -0.001 | 1.00 | -0.011 | 0.014 |
| JEE Main score | 79.74 | 38.53 | -0.662 | 0.587 |
| Took JEE Main | 0.66 | 0.48 | 0.001 | 0.009 |

Notes: Estimates are calculated using the second cohort of students. Means and standard deviations for general category faculty characteristics are reported in Panel A. Means and standard deviations for all sampled students are reported in Panel B. The sample of colleges with random assignment (12 colleges) is used, and the unit of analysis is a student-course. The data capture 2289 students and 650 faculty. The reservation vs general category differences control for course fixed effects, and corresponding standard errors are clustered at the faculty level. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table G2: Regressions for Additional Educational Outcomes, Reservation Category Faculty Interacted with Reservation Category Students (Second cohort of students)

| | I EE test (Year 4) | II CS test (Year 4) | III Expected Graduation (Year 4) | IV Plans for Graduate School (Year 4) |
|------------------------------------|--------------------------|---------------------------|--|---|
| Res. Cat. Faculty | -0.041 (0.033) | -0.001 (0.049) | -0.00008 (0.00011) | 0.010 (0.013) |
| Res. Cat. Student | -0.300** (0.136) | -0.120 (0.137) | -0.00028 (0.00047) | 0.030 (0.051) |
| R.C. Faculty \times R.C. Student | 0.025 (0.026) | -0.051 (0.032) | 0.00002 (0.00008) | -0.008 (0.011) |
| Faculty controls | Yes | Yes | Yes | Yes |
| Student controls | Yes | Yes | Yes | Yes |
| Mean | 0.00 | -0.03 | 0.99 | 0.51 |
| N | 1060 | 510 | 2247 | 2083 |

Notes: The dependent variables are measured at the end of year 4 and are (I) standardized test score for the electrical engineering (EE) test, (II) standardized test score for the computer science (CS) test, (III) whether the student expected to graduate, and (IV) whether the student aspired for graduate school after completing their program. Res. Cat. faculty is the percentage of reservation category faculty who taught courses taken by the student, and is rescaled to capture the effect of changing the reservation category faculty percentage by 10 percentage points (e.g. from 0.50 to 0.60). Student controls include gender, age, and parents' education. Faculty controls include reservation category status, professor rank, experience, highest degree, elite college, and gender. The coefficients from Specification III are the marginal effects from a probit model between the expected graduation (0/1) variable and the listed covariates. All models are run on the second cohort of students for the sample with random assignment (12 colleges), where each observation is a student. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

H Extra Measures of Research

Table H1: Regressions for Number of SCI, EI or SSCI Publications per Year, Reservation Category vs. General Category Faculty using the National Sample

| | I | II | III | IV |
|--------------------------------|---------------------|-----------------------|-----------------------|-----------------------|
| Res. Cat. Faculty | -0.0541 (0.0427) | 0.0036 (0.0417) | 0.0037 (0.0413) | -0.0031 (0.0407) |
| Associate professor | | 0.2106*** (0.0744) | 0.1301* (0.0681) | 0.1336** (0.0667) |
| Professor | | 0.9432*** (0.2142) | 0.7717*** (0.2141) | 0.7622*** (0.2134) |
| Experience in years | | 0.0106** (0.0053) | 0.0081 (0.0054) | 0.0078 (0.0053) |
| Highest degree PhD | | | 0.2517*** (0.0760) | 0.2554*** (0.0743) |
| Highest degree PhD in progress | | | 0.0084 (0.0368) | 0.0061 (0.0368) |
| Degree from elite college | | | 0.0874* (0.0527) | 0.0844 (0.0522) |
| Female | | | | -0.0648** (0.0294) |
| Mean | 0.53 | 0.53 | 0.53 | 0.53 |
| N | 2691 | 2685 | 2680 | 2679 |

The dependent variables are the number of articles authored by a faculty that were published in SCI (Science Citation Index), SSCI (Social Sciences Citation Index), and EI (Engineering Index) listed journals. The regressions use department-level sampling weights, and are run at the faculty level for the national sample (50 colleges). All specifications include college and department fixed effects. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table H2: Regressions for Government Funding Received, Reservation vs. General Category Faculty using the National Sample

| | I | II | III | IV |
|--------------------------------|---------------------|-----------------------|-----------------------|-----------------------|
| Res. Cat. Faculty | -0.0059 (0.0090) | 0.0026 (0.0085) | 0.0018 (0.0086) | 0.0010 (0.0084) |
| Associate professor | | 0.0253 (0.0159) | 0.0164 (0.0172) | 0.0168 (0.0171) |
| Professor | | 0.1536*** (0.0352) | 0.1353*** (0.0358) | 0.1341*** (0.0358) |
| Experience in years | | 0.0010 (0.0009) | 0.0011 (0.0009) | 0.0010 (0.0009) |
| Highest degree PhD | | | 0.0212 (0.0169) | 0.0217 (0.0168) |
| Highest degree PhD in progress | | | -0.0184** (0.0093) | -0.0186** (0.0093) |
| Degree from elite college | | | -0.0120 (0.0117) | -0.0124 (0.0117) |
| Female | | | | -0.0077 (0.0080) |
| Mean | 0.097 | 0.097 | 0.097 | 0.097 |
| N | 2691 | 2685 | 2680 | 2679 |

The dependent variable is government research funding received at college (0/1). The regressions use department-level sampling weights, and are run at the faculty level for the national sample (50 colleges). All specifications include college and department fixed effects. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table H3: Regressions for Private Funding Received, Reservation vs. General Category Faculty using the National Sample

| | I | II | III | IV |
|--------------------------------|---------------------|---------------------|----------------------|------------------------|
| Res. Cat. Faculty | -0.0021 (0.0074) | -0.0006 (0.0077) | -0.0003 (0.0077) | -0.0024 (0.0078) |
| Associate professor | | -0.0039 (0.0100) | -0.0131 (0.0112) | -0.0121 (0.0112) |
| Professor | | 0.0216 (0.0211) | 0.0036 (0.0222) | 0.0006 (0.0222) |
| Experience in years | | 0.0007 (0.0007) | 0.0005 (0.0007) | 0.0004 (0.0007) |
| Highest degree PhD | | | 0.0323** (0.0143) | 0.0334** (0.0143) |
| Highest degree PhD in progress | | | 0.0128 (0.0086) | 0.0121 (0.0086) |
| Degree from elite college | | | -0.0105* (0.0062) | -0.0114* (0.0062) |
| Female | | | | -0.0200*** (0.0062) |
| Mean | 0.029 | 0.029 | 0.029 | 0.029 |
| N | 2691 | 2685 | 2680 | 2679 |

The dependent variable is private research funding received at college (0/1). The regressions use department-level sampling weights, and are run at the faculty level for the national sample (50 colleges). All specifications include college and department fixed effects. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table H4: Regressions for Research, Funding, and Administration for the Sample of Colleges with Random Assignment, Reservation vs. General Category Faculty (Student-Course Level)

| | I Publications | II International Publications | III SCI/EI/SSCI Publications | IV Funding Received | V Administrative Position |
|--------------------------------|-------------------|-------------------------------------|------------------------------------|---------------------------|---------------------------------|
| Res. Cat. Faculty | -0.37 (0.27) | -0.29** (0.14) | -0.10 (0.09) | -0.03 (0.02) | 0.03 (0.05) |
| Associate professor | 1.11*** (0.37) | 0.92*** (0.20) | 0.33** (0.16) | 0.07 (0.07) | 0.01 (0.07) |
| Professor | 1.80*** (0.69) | 1.16*** (0.45) | 0.72** (0.31) | -0.01 (0.07) | 0.02 (0.11) |
| Experience in years | -0.06** (0.03) | -0.02* (0.01) | -0.02* (0.01) | 0.00 (0.00) | 0.02*** (0.00) |
| Highest degree PhD | 2.37*** (0.78) | 0.92*** (0.29) | 0.66** (0.26) | 0.23*** (0.07) | 0.19* (0.11) |
| Highest degree PhD in progress | 1.06*** (0.39) | 0.53*** (0.17) | 0.12 (0.11) | 0.01 (0.03) | -0.11* (0.06) |
| Degree from elite college | 0.16 (0.32) | 0.00 (0.18) | 0.00 (0.12) | -0.09*** (0.04) | 0.00 (0.06) |
| Female | -0.24 (0.24) | -0.20 (0.12) | -0.07 (0.09) | 0.09*** (0.03) | 0.06 (0.05) |
| N | 37970 | 37970 | 37970 | 37970 | 37970 |

Notes: Dependent variables refer to annual publications (I), annual international publications (II), annual international SCI/EI/SSCI publications (III), funding received (IV), and administrative position held (V). The regressions are run at the student-course level for the sample of colleges with random assignment (12 colleges). All specifications include student fixed effects and course fixed effects. Standard errors are clustered at the faculty level. Significance levels: *** $p < 0.01$; ** $p < 0.05$; * $p < 0.1$.